GAUSSIAN MIXTURE PENALIZATION FOR TRAJECTORY OPTIMIZATION PROBLEMS

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CMAP Ecole Polytechnique - INRIA¹ Safety Line²

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• 20 000 airplanes — 80 000 flights per day,

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- 20 000 airplanes 80 000 flights per day,
- Should double until 2033,
- Responsible for 3% of CO₂ emissions,
- Accounts for 30% of operational cost for an airline,
- Rectilinear climb trajectories at full thrust.



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Optimal Control Problem

s.t.
$$\begin{cases} \min_{(\mathbf{x},\mathbf{u})\in\mathbb{X}\times\mathbb{U}} \int_{0}^{t_{f}} C(t,\mathbf{u}(t),\mathbf{x}(t))dt, \\ \dot{\mathbf{x}} = g(t,\mathbf{u},\mathbf{x}), & \text{for a.e. } t \in [0,t_{f}], \\ \Phi(\mathbf{x}(0),\mathbf{x}(t_{f})) \in K_{\Phi}, \\ \mathbf{u}(t) \in U_{ad}, \quad \mathbf{x}(t) \in X_{ad}, & \text{for a.e. } t \in [0,t_{f}], \\ c(\mathbf{u}(t),\mathbf{x}(t)) \leq 0, & \text{for all } t \in [0,t_{f}]. \end{cases}$$
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See e.g. [Rommel et al., 2017a] and [Rommel et al., 2017b]

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Is \hat{z} inside the validity region of the dynamics model \hat{g} ? Does it look like a real aicraft trajectory?

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Pilots acceptance

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Pilots acceptance



Air Traffic Control¹

¹NATS UK air traffic control

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Pilots acceptance Air Traffic Control¹ How can we quantify the closeness from the optimized trajectory to the set of real flights?

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Let X be a random variable following an absolutely continuous probability distribution with density function f depending on a parameter θ . Then the function

$$\mathcal{L}(\theta|x) = f_{\theta}(x) \tag{1}$$

considered as a function of θ , is the likelihood function of *theta*, given the outcome x of X.

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In our case:

- the optimized trajectory plays the role of θ,
- the set of real flights plays the role of *x*,

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Assumption: We suppose that the real flights are observations of the same functional random variable $Z = (Z_t)$ valued in $\mathcal{C}(\mathbb{T}, E)$, with E compact subset of \mathbb{R}^d and $\mathbb{T} = [0, t_f]$.

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Problem: Computation of probability densities in infinite dimensional space is untractable...

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 Standard approach FDA: use FPCA to decompose the data in a small number of coefficients



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- Standard approach FDA: use FPCA to decompose the data in a small number of coefficients
- Or: we can aggregate the marginal densities



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WHY DOES IT MAKE SENSE FOR THIS TYPE OF DATA?

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WHY DOES IT MAKE SENSE FOR THIS TYPE OF DATA? Likely values of flight variables during climb are strongly dependent on the altitude

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How do we aggregate the marginal LIKELIHOODS?

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How do we aggregate the marginal likelihoods?

• f_t marginal density of Z, i.e. probability density function of Z_t ,

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Why not average over time ?...

$$\frac{1}{t_f}\int_0^{t_f}f_t(\boldsymbol{y}(t))dt$$

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Marginal densities may have really different shapes

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- f_t marginal density of Z, i.e. probability density function of Z_t ,
- y new trajectory,
- $f_t(\mathbf{y}(t))$ marginal likelihood of \mathbf{y} at t, i.e. likelihood of observing $Z_t = \mathbf{y}(t)$.

Mean marginal likelihood [Rommel et al., 2018]

$$\mathsf{MML}(Z, \mathbf{y}) = \frac{1}{t_f} \int_0^{t_f} \psi[f_t, \mathbf{y}(t)] dt,$$

where $\psi : L^1(E, \mathbb{R}_+) \times \mathbb{R} \to [0; 1]$ is a continuous scaling map.

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Possible scalings are the normalized density

$$\psi[f_t, \boldsymbol{y}(t)] := rac{\boldsymbol{y}(t)}{\displaystyle\max_{z\in E} f_t(z)},$$

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How do we aggregate the marginal LIKELIHOODS?

Possible scalings are the normalized density

$$\psi[f_t, \boldsymbol{y}(t)] := rac{\boldsymbol{y}(t)}{\displaystyle\max_{z \in E} f_t(z)},$$

or the confidence level



$$\psi[f_t, \boldsymbol{y}(t)] := \mathbb{P}\left(f_t(Z_t) \leq f_t(\boldsymbol{y}(t))\right).$$

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How do we deal with sampled curves?

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How do we deal with sampled curves?

In practice, the m trajectories are sampled at variable discrete times:

$$\begin{aligned} \mathcal{T}^{D} &:= \{ (t_{j}^{r}, z_{j}^{r}) \}_{\substack{1 \leq j \leq n \\ 1 \leq r \leq m}} \subset \mathbb{T} \times E, \qquad \qquad z_{j}^{r} &:= \mathbf{z}^{r}(t_{j}^{r}), \\ \mathcal{Y} &:= \{ (\tilde{t}_{j}, y_{j}) \}_{j=1}^{\tilde{n}} \subset \mathbb{T} \times E, \qquad \qquad y_{j} &:= \mathbf{y}(\tilde{t}_{j}). \end{aligned}$$

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How do we deal with sampled curves?

In practice, the *m* trajectories are sampled at variable discrete times:

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Hence, we approximate the MML using a Riemann sum which aggregates consistent estimators $\hat{f}_{\tilde{t}_i}^m$ of the marginal densities $f_{\tilde{t}_j}$:

$$\mathsf{EMML}_m(\mathcal{T}^D,\mathcal{Y}) := rac{1}{t_f} \sum_{j=1}^{\tilde{n}} \psi[\hat{f}^m_{\tilde{t}_j},y_j] \Delta \tilde{t}_j.$$

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Suppose that sampling times {t_j^r : j = 1,..., n; r = 1,..., m} are i.i.d. sampled from r.v. T, indep. Z;

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 \Rightarrow We could apply SOA conditional density estimation techniques, such as LS-CDE [Sugiyama et al., 2010],

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 \Rightarrow We could apply SOA conditional density estimation techniques, such as LS-CDE [Sugiyama et al., 2010],

 \Rightarrow Instead, we choose to use a fine partitioning of the time domain.

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PARTITION BASED MARGINAL DENSITY ESTIMATION



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PARTITION BASED MARGINAL DENSITY ESTIMATION



Idea: to average in time the marginal densities over small bins by applying classical multivariate density estimation techniques to each subset.

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Assumption 1 - Positive time density $\nu \in L^{\infty}(E, \mathbb{R}_+)$ density function of *T*, s.t.

$$u_+ := \mathop{\mathrm{ess\,sup}}_{t\in\mathbb{T}}
u(t) < \infty, \qquad
u_- := \mathop{\mathrm{ess\,inf}}_{t\in\mathbb{T}}
u(t) > 0.$$

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ASSUMTION 2 - LIPSCHITZ IN TIME Function $(t, z) \in \mathbb{T} \times E \mapsto f_t(z)$ is continuous and $|f_{t_1}(z) - f_{t_2}(z)| \le L|t_1 - t_2|, \qquad L > 0.$

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Assumption 3 - Shrinking Bins The homogeneous partition $\{B_{\ell}^m\}_{\ell=1}^{q_m}$ of [0; t_f], with binsize b_m , is s.t.

$$\lim_{m\to\infty}b_m=0,\qquad \lim_{m\to\infty}mb_m=\infty.$$

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Assumption 4 - I.I.D. Consistency

•
$$\mathcal{S} = \{(z_k)_{k=1}^N \in E^N : N \in \mathbb{N}^*\}$$
 set of finite sequences,

- $\Theta: \mathcal{S} \to L^1(E, \mathbb{R}_+)$ multivariate density estimation statistic,
- ${\cal G}$ arbitrary family of probability density functions on ${\cal E}$, $ho \in {\cal G}$,
- S_{ρ}^{N} <u>i.i.d</u> sample of size N drawn from ρ valued in S.

The estimator obtained by applying Θ to S_{ρ}^{N} , denoted by

$$\hat{\rho}^{\mathsf{N}} := \Theta[S^{\mathsf{N}}_{\rho}] \in L^1(E, \mathbb{R}_+),$$

is a (pointwise) consistent density estimator, uniformly in ρ :

For all $z \in E, \varepsilon > 0, \alpha_1 > 0$, there is $N_{\varepsilon,\alpha_1} > 0$ such that, for any $\rho \in \mathcal{G}$, $N \ge N_{\varepsilon,\alpha_1} \Rightarrow \mathbb{P}\left(\left|\hat{\rho}^N(z) - \rho(z)\right| < \varepsilon\right) > 1 - \alpha_1.$

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We denote by:

- $\ell^m(t) := \left\lceil \frac{t}{b_m} \right\rceil$ maps time to index of bin containing it;
- *f*^m_{ℓ^m(t)} := Θ[*T*^m_{ℓ^m(t)}] estimator trained using subset of data points *T*^m_{ℓ^m(t)}
 whose sampling times fall in the bin containing t;

THEOREM 1 - [ROMMEL ET AL., 2018]

Under assumptions 1 to 4, for any $z \in E$ and $t \in \mathbb{T}$, $\hat{f}^m_{\ell^m(t)}(z)$ consistently approximates the marginal density $f_t(z)$ as the number of curves *m* grows:

$$orall arepsilon > 0, \quad \lim_{m o \infty} \mathbb{P}\left(|\hat{f}^m_{\ell^m(t)}(z) - f_t(z)| < arepsilon
ight) = 1.$$

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MARGINAL DENSITY ESTIMATION RESULTS



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MARGINAL DENSITY ESTIMATION RESULTS



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MARGINAL DENSITY ESTIMATION RESULTS



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• Training set of m = 424 flights $\simeq 334$ 531 point observations,

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- Training set of m = 424 flights $\simeq 334$ 531 point observations,
- Test set of 150 flights

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- Training set of m = 424 flights $\simeq 334$ 531 point observations,
- Test set of 150 flights = 50 real flights (*Real*), 50 optimized flights with operational constraints (*Opt1*) and 50 optimized flights without constraints (*Opt2*);

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- Discrimination power comparison with (gmm-)FPCA and (integrated) LS-CDE:

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- Discrimination power comparison with (gmm-)FPCA and (integrated) LS-CDE:

VAR.	Estimated Likelihoods		
	Real	Opt1	Opt2
MML	$\textbf{0.63} \pm \textbf{0.07}$	$\textbf{0.43}\pm\textbf{0.08}$	$\textbf{0.13} \pm \textbf{0.02}$
FPCA	0.16 ± 0.12	$6.4\text{E-}03 \pm 3.8\text{E-}03$	$3.6\text{E-}03 \pm 5.4\text{E-}03$
LS-CDE	0.77 ± 0.05	0.68 ± 0.04	0.49 ± 0.06

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$MML \ {\tt penalty}$

The MML can be used not only to assess the optimization solutions, but also to penalize the optimization itself:

$$\min_{\substack{(\mathbf{x}, \mathbf{u}) \in \mathbb{X} \times \mathbb{U} \\ \text{s.t.}}} \int_{0}^{t_{f}} C(t, \mathbf{u}(t), \mathbf{x}(t)) dt$$
$$\begin{cases} \dot{\mathbf{x}} = g(t, \mathbf{u}, \mathbf{x}), & \text{for a.e. } t \in [0, t_{f}], \\ \Phi(\mathbf{x}(0), \mathbf{x}(t_{f})) \in K_{\Phi}, \\ \mathbf{u}(t) \in U_{ad}, \quad \mathbf{x}(t) \in X_{ad}, & \text{for a.e. } t \in [0, t_{f}], \\ c(\mathbf{u}(t), \mathbf{x}(t)) \leq 0, & \text{for all } t \in [0, t_{f}]. \end{cases}$$
(OCP)

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$$\sup_{\substack{(\mathbf{x}, \mathbf{u}) \in \mathbb{X} \times \mathbb{U} \\ \Phi(\mathbf{x}(0), \mathbf{x}(t_{f})) \in K_{\Phi}, \\ \mathbf{u}(t) \in U_{ad}, \quad \mathbf{x}(t) \in X_{ad}, \quad \text{for a.e. } t \in [0, t_{f}],$$

$$C(\mathbf{u}(t), \mathbf{x}(t)) \leq 0, \quad \text{for all } t \in [0, t_{f}].$$

$$(OCP)$$

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 $\bullet~\lambda$ sets trade-off between a fuel minimization and a likelihood maximization,

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(OCP)

- $\bullet~\lambda$ sets trade-off between a fuel minimization and a likelihood maximization,
- If (OCP) is solved using NLP techniques, parametric estimator of MML is needed.

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GAUSSIAN MIXTURE MODEL FOR MARGINAL DENSITIES

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GAUSSIAN MIXTURE MODEL FOR MARGINAL DENSITIES

$$f_{t}(z) = \sum_{k=1}^{K} w_{t,k} \phi(z, \mu_{t,k}, \Sigma_{t,k}),$$

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$$\sum_{k=1}^{K} w_{t,k} = 1, \qquad w_{t,k} \ge 0,$$

$$\phi(z, \mu, \Sigma) := \frac{1}{\sqrt{(2\pi)^{d} \det \Sigma}} e^{-\frac{1}{2}(z-\mu)^{\top} \Sigma^{-1}(z-\mu)}.$$

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$$\phi(z, \mu, \Sigma) := \frac{1}{\sqrt{(2\pi)^{d} \det \Sigma}} e^{-\frac{1}{2}(z-\mu)^{\top} \Sigma^{-1}(z-\mu)}.$$

Assuming that the number of components is known, the weights $w_{t,k}$, means $\mu_{t,k}$ and covariance matrices $\Sigma_{t,k}$ need to be estimated.

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For K = 1, maximum likelihood estimates have closed form:

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For K = 1, maximum likelihood estimates have closed form:

$$\mathcal{L}(\mu_{t,1}, \Sigma_{t,1} | z_1, \dots, z_N) = \prod_{i=1}^N \frac{1}{\sqrt{(2\pi)^d \det \Sigma_{t,1}}} e^{-\frac{1}{2}(z-\mu_{t,1})^\top \Sigma_{t,1}^{-1}(z-\mu_{t,1})}$$

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$$\hat{\theta} := (\hat{\mu}_{t,1}, \hat{\Sigma}_{t,1}) = \arg\min_{(\mu_{t,1}, \Sigma_{t,1})} \sum_{i=1}^{N} \left(\log \det \Sigma_{t,1} + (z_i - \mu_{t,1})^{\top} \Sigma_{t,1}^{-1} (z_i - \mu_{t,1}) \right)$$

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$$\hat{\mu}_{t,1} = \frac{1}{N} \sum_{i=1}^{N} z_i, \qquad \hat{\Sigma}_{t,1} = \frac{1}{N} \sum_{i=1}^{N} (z_i - \hat{\mu}_{t,1}) (z_i - \hat{\mu}_{t,1})^{\top}.$$

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GMM PENALIZATION FOR OCP

• Hidden random variable J valued on $\{1, \ldots, K\}$,

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GMM PENALIZATION FOR OCP

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- If i^{th} observation $J_i = k$, then z_i was drawn from the k^{th} component,

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GMM PENALIZATION FOR OCP

- Hidden random variable J valued on $\{1, \ldots, K\}$,
- If i^{th} observation $J_i = k$, then z_i was drawn from the k^{th} component,
- Group observations by component and compute $(\hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})$ with K = 1 maximum likelihood formulas.

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- If i^{th} observation $J_i = k$, then z_i was drawn from the k^{th} component,
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EXPECTATION-MAXIMIZATION - [DEMPSTER ET AL., 1977] Initialization: $\hat{\theta} = (\hat{w}_{t,k}, \hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})_{k=1}^{K} = (w_{t,k}^{0}, \mu_{t,k}^{0}, \Sigma_{t,k}^{0})_{k=1}^{K}$,

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$$\begin{vmatrix} \hat{w}_{t,k} = \frac{1}{N} \sum_{i=1}^{N} \hat{\pi}_{k,i}, \\ \hat{\pi}_{k,i} := \mathbb{P}(J_i = k | \hat{\theta}_t, Z_h) = \frac{\hat{\mu}_{t,k} \phi(z_i, \hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})}{\sum_{j=1}^{N} \hat{w}_{t,k} \phi(z_j, \hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})}. \end{vmatrix}$$

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Maximization:

$$\widehat{\mu}_{t,k} = \frac{\sum_{i=1}^{N} \widehat{\pi}_{k,i} z_i}{\sum_{i=1}^{N} \widehat{\pi}_{k,i}}, \qquad \widehat{\Sigma}_{t,k} = \frac{\sum_{i=1}^{N} \widehat{\pi}_{k,i} (z_i - \widehat{\mu}_{t,k}) (z_i - \widehat{\mu}_{t,k})^{\top}}{\sum_{i=1}^{N} \widehat{\pi}_{k,i}}.$$

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PENALTY EFFECT



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CONSUMPTION X ACCEPTABILITY TRADE-OFF



FIGURE: Average over 20 flights of the fuel consumption and MML score (called acceptability here) of optimized trajectories with varying MML-penalty weight λ .

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General probabilistic criterion for quantifying the closeness between a curve and a set random trajectories,

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- General probabilistic criterion for quantifying the closeness between a curve and a set random trajectories,
- ② Class of consistent plug-in estimators, based on "histogram" of multivariate density estimators,

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- General probabilistic criterion for quantifying the closeness between a curve and a set random trajectories,
- Class of consistent plug-in estimators, based on "histogram" of 2 multivariate density estimators,
- Applicable to the case of aircraft climb trajectories,
 - Competitive with other well-established SOA approaches,

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- \Rightarrow How could we automatically set the trade-off ?...

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THANK YOU FOR YOUR ATTENTION !!

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