

EXPLORATION DE DONNÉES POUR L'OPTIMISATION DE TRAJECTOIRES AÉRIENNES

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Encadrant Safety Line: Baptiste Gregorutti

Soutenance de thèse, 26 octobre 2018



MOTIVATION

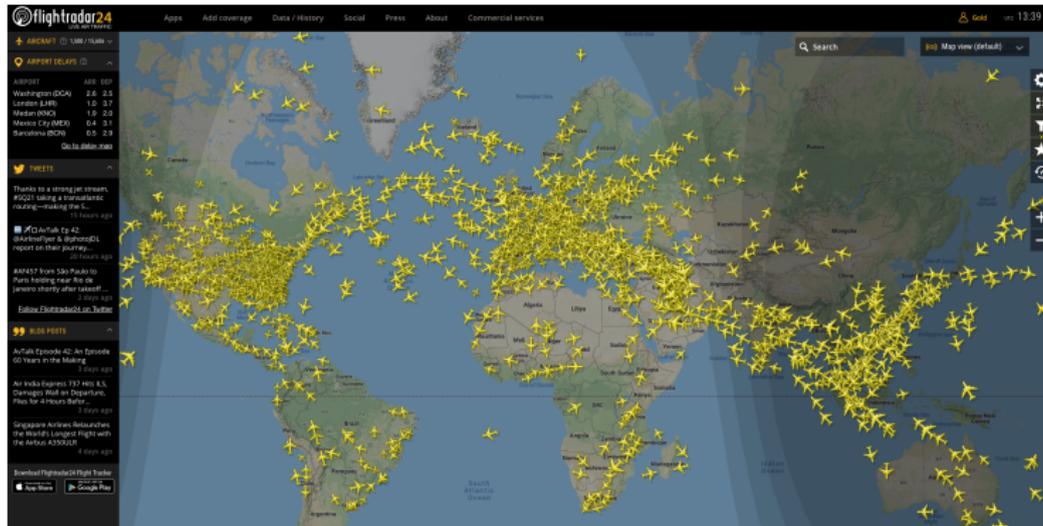


FIGURE: World air traffic - source: www.flightradar24.com

MOTIVATION

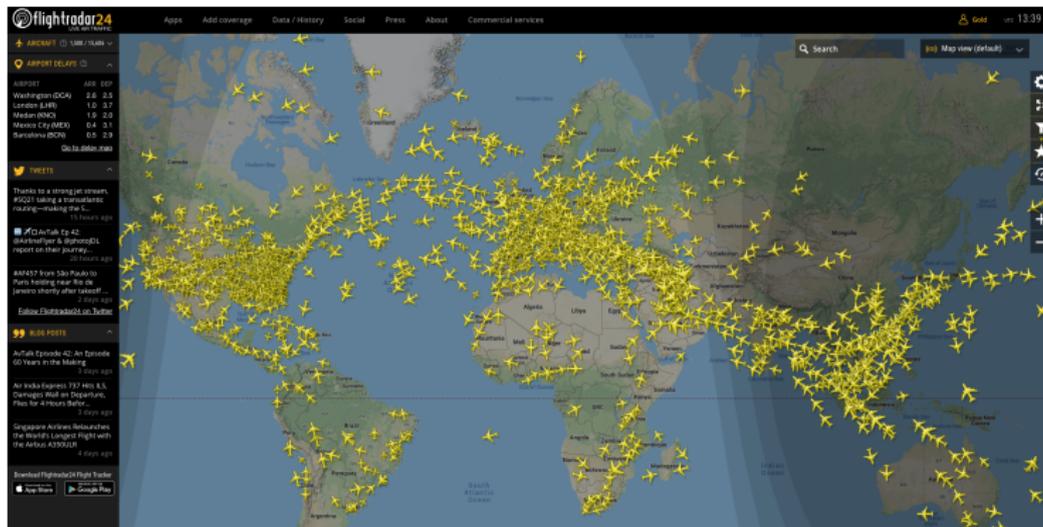


FIGURE: World air traffic - source: www.flightradar24.com

- 20 000 airplanes — 80 000 flights per day,

MOTIVATION

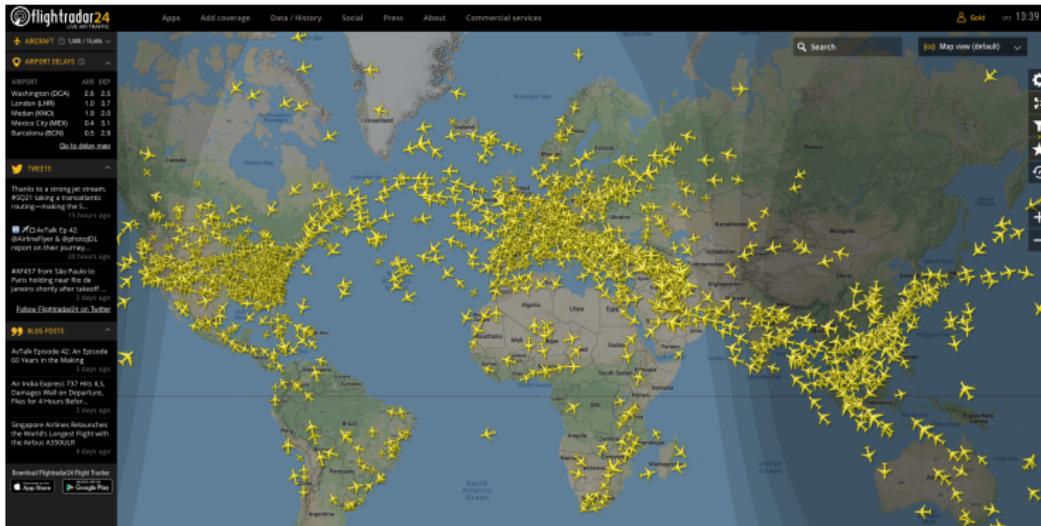


FIGURE: World air traffic - source: www.flightradar24.com

- 20 000 airplanes — 80 000 flights per day,
- Should double until 2033,

MOTIVATION

How to tackle this problem ?

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- 1 New hardware ?

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- 1 New hardware ?
- 2 **Better use of existing fleet,**

MOTIVATION

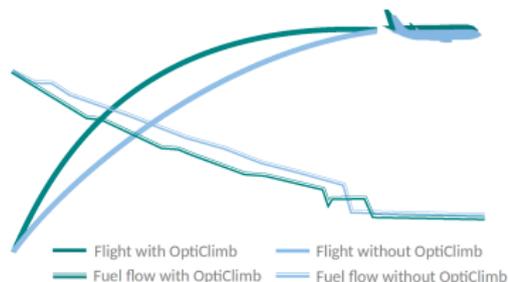
How to tackle this problem ?

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 - Climb is the most consuming flight phase...

MOTIVATION

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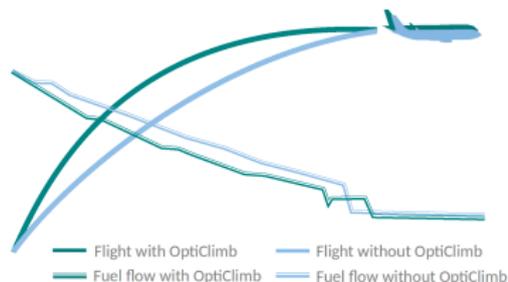
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MOTIVATION

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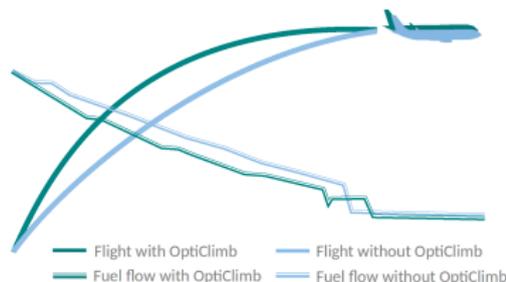
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 - Thousands of variables recorded every second,



MOTIVATION

How to tackle this problem ?

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 **OPTI CLIMB**
More savings, less CO₂

OPTICLIMB

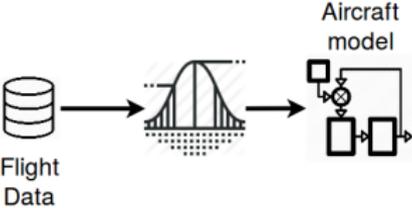


Flight
Data

Time

Many days before flight...

OPTICLIMB

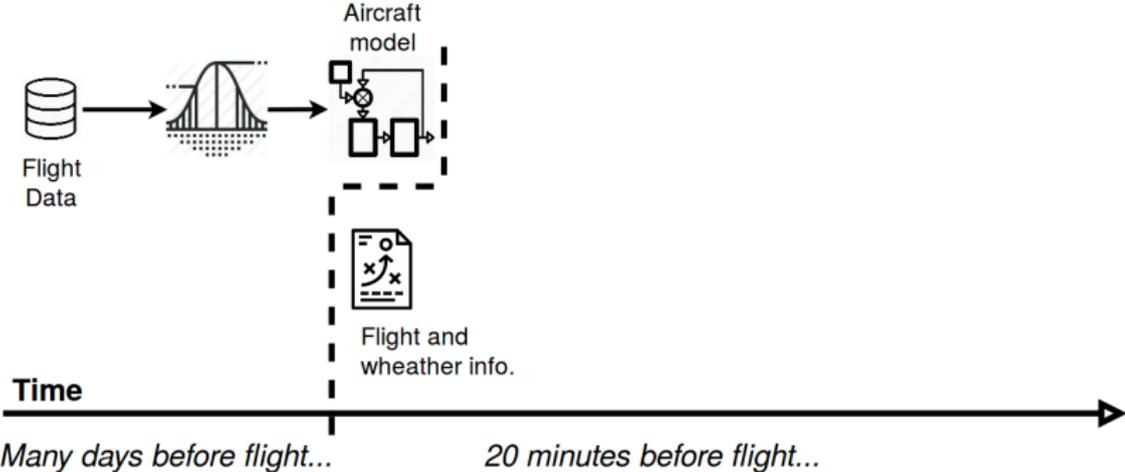


Time

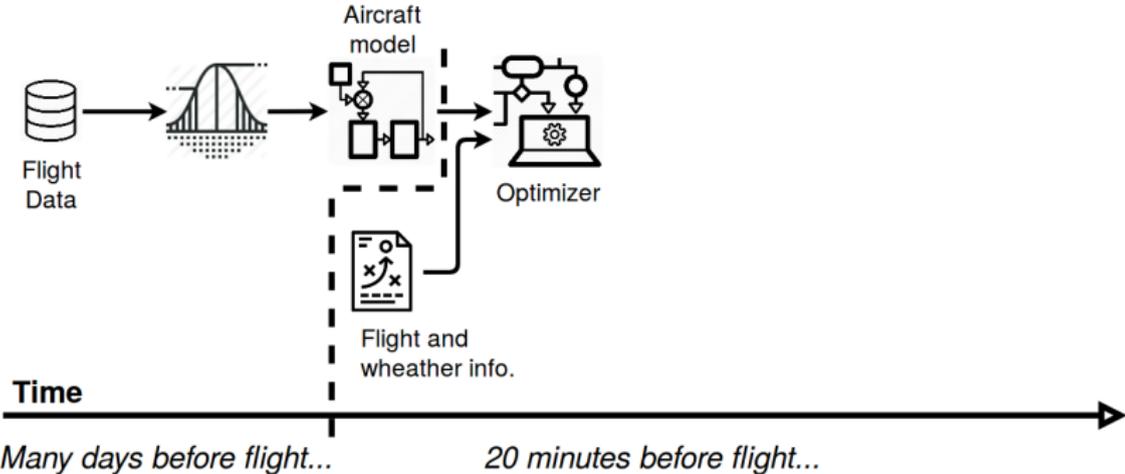


Many days before flight...

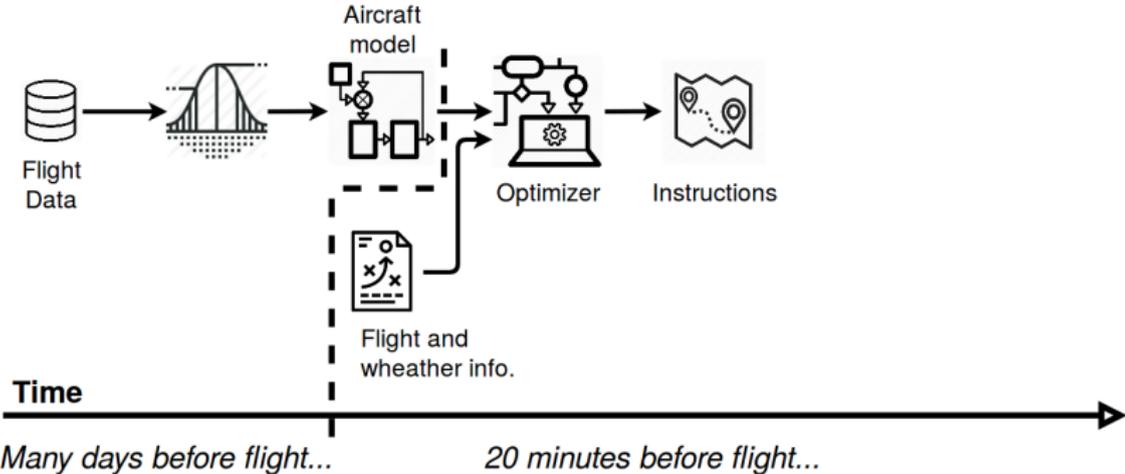
OPTICLIMB



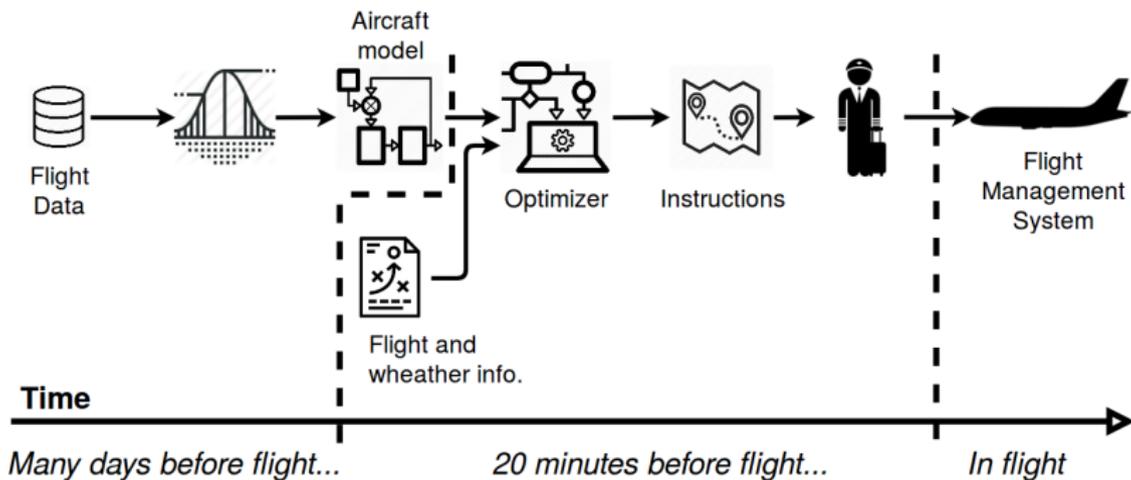
OPTICLIMB



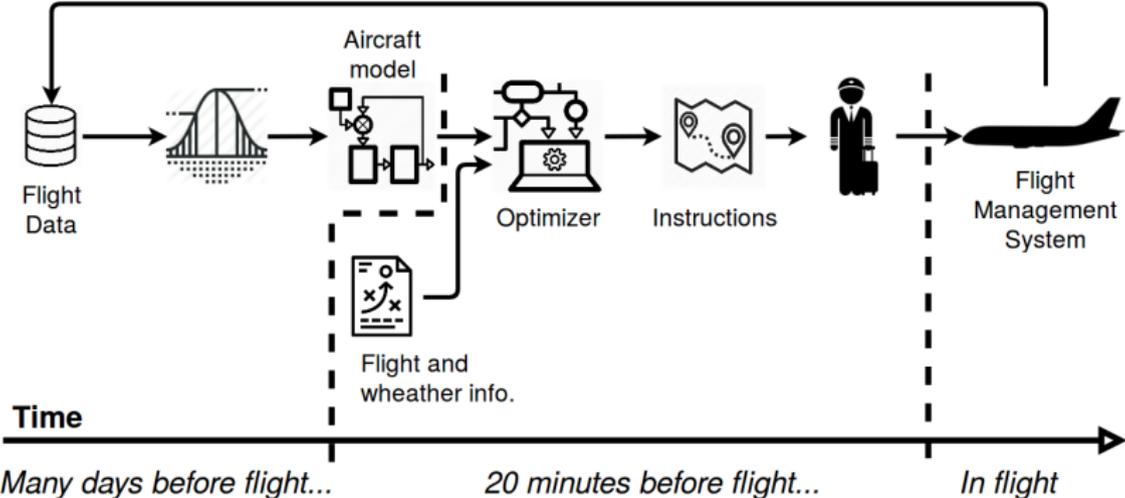
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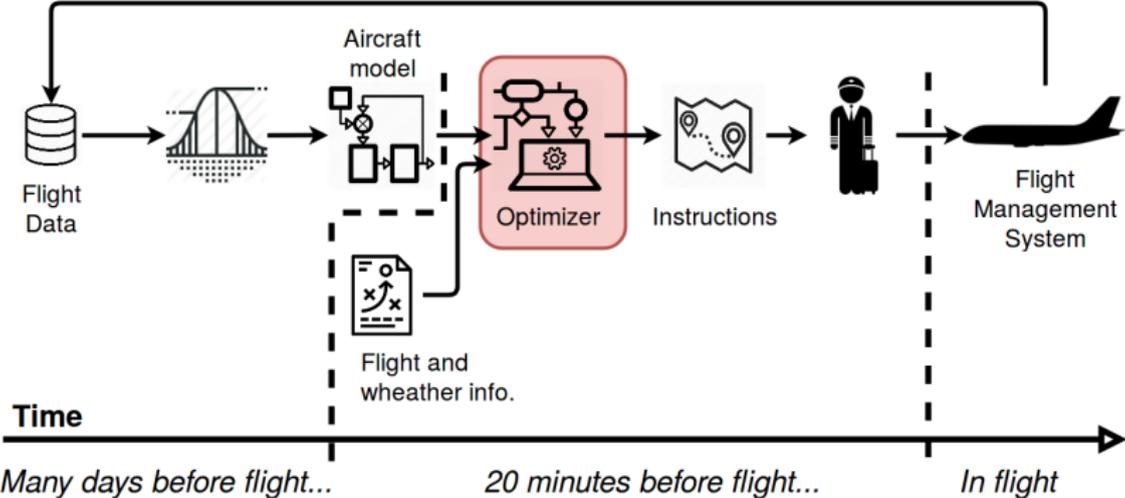
OPTICLIMB



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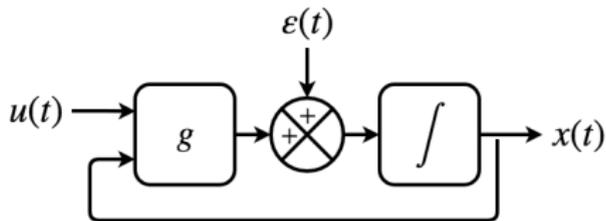
OPTICLIMB



TRAJECTORY OPTIMIZATION

Dynamics:

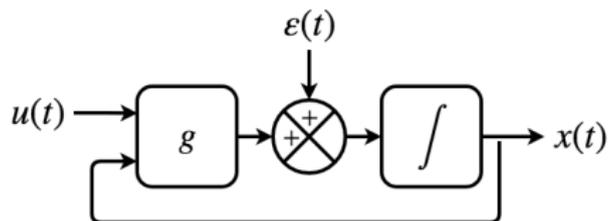
$$\dot{\mathbf{x}}(t) = \mathbf{g}(\mathbf{u}(t), \mathbf{x}(t)) + \boldsymbol{\varepsilon}(t)$$



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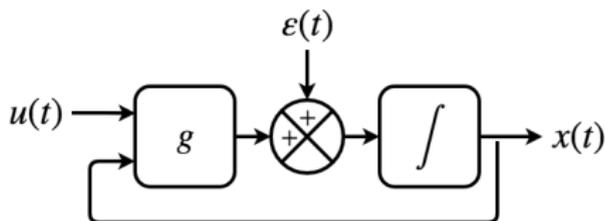


Optimization objective: $\int_0^{t_f} C(\mathbf{u}(t), \mathbf{x}(t)) dt$

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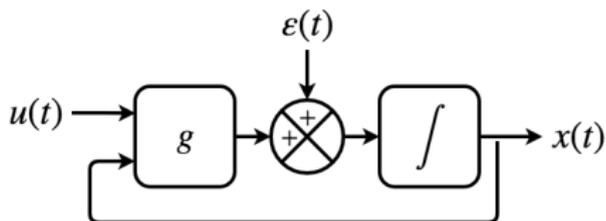


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TRAJECTORY OPTIMIZATION

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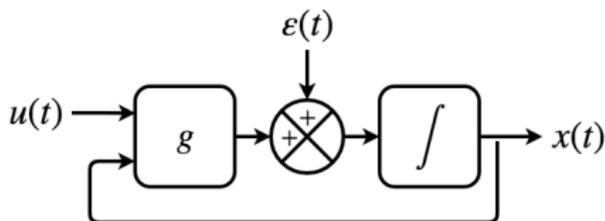


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TRAJECTORY OPTIMIZATION

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Optimization objective: $\int_0^{t_f} C(\mathbf{u}(t), \mathbf{x}(t)) dt \leftarrow$ , , , 

Flight constraints:

$$\begin{cases} \Phi(\mathbf{x}(0), \mathbf{x}(t_f)) \in K_\Phi \\ \mathbf{u}(t) \in U_{ad}, \quad \mathbf{x}(t) \in X_{ad}, \\ c(\mathbf{u}(t), \mathbf{x}(t)) \leq 0, \end{cases}$$

Initial and final conditions

Flight domain

Operational path constraints

TRAJECTORY OPTIMIZATION

OPTIMAL CONTROL PROBLEM

$$\begin{aligned} & \min_{(\mathbf{x}, \mathbf{u}) \in \mathcal{X} \times \mathcal{U}} \int_0^{t_f} C(\mathbf{u}(t), \mathbf{x}(t)) dt, \\ \text{s.t. } & \begin{cases} \dot{\mathbf{x}}(t) = \mathbf{g}(\mathbf{u}(t), \mathbf{x}(t)) + \varepsilon(t), & \text{a.e. } t \in [0, t_f], \\ \Phi(\mathbf{x}(0), \mathbf{x}(t_f)) \in K_\Phi, \\ \mathbf{u}(t) \in U_{ad}, \quad \mathbf{x}(t) \in X_{ad}, & \text{a.e. } t \in [0, t_f], \\ c(\mathbf{u}(t), \mathbf{x}(t)) \leq 0, & \text{a.e. } t \in [0, t_f]. \end{cases} \end{aligned} \quad (\text{OCP})$$

TRAJECTORY OPTIMIZATION

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TRAJECTORY OPTIMIZATION

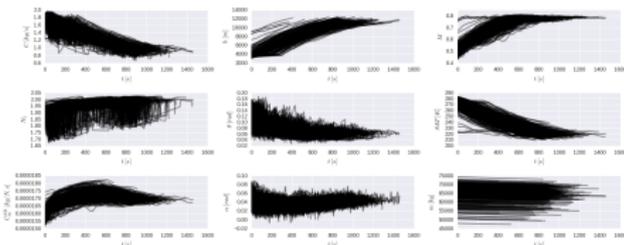
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SYSTEM IDENTIFICATION



Black box



QAR data

TRAJECTORY OPTIMIZATION

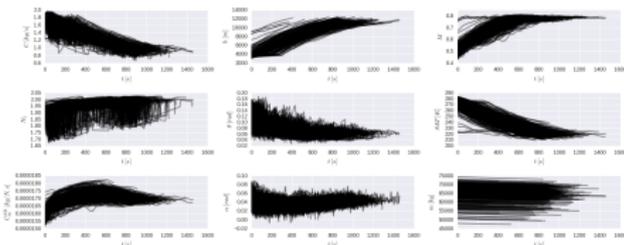
APPROXIMATE OPTIMAL CONTROL PROBLEM

$$\min_{(\mathbf{x}, \mathbf{u}) \in \mathbb{X} \times \mathbb{U}} \int_0^{t_f} C(\mathbf{u}(t), \mathbf{x}(t)) dt,$$
$$\text{s.t. } \begin{cases} \dot{\mathbf{x}}(t) = \hat{\mathbf{g}}(\mathbf{u}(t), \mathbf{x}(t)), & \text{a.e. } t \in [0, t_f], \\ \Phi(\mathbf{x}(0), \mathbf{x}(t_f)) \in K_\Phi, & \\ \mathbf{u}(t) \in U_{ad}, \quad \mathbf{x}(t) \in X_{ad}, & \text{a.e. } t \in [0, t_f], \\ c(\mathbf{u}(t), \mathbf{x}(t)) \leq 0, & \text{a.e. } t \in [0, t_f]. \end{cases} \quad (\text{A OCP})$$

SYSTEM IDENTIFICATION

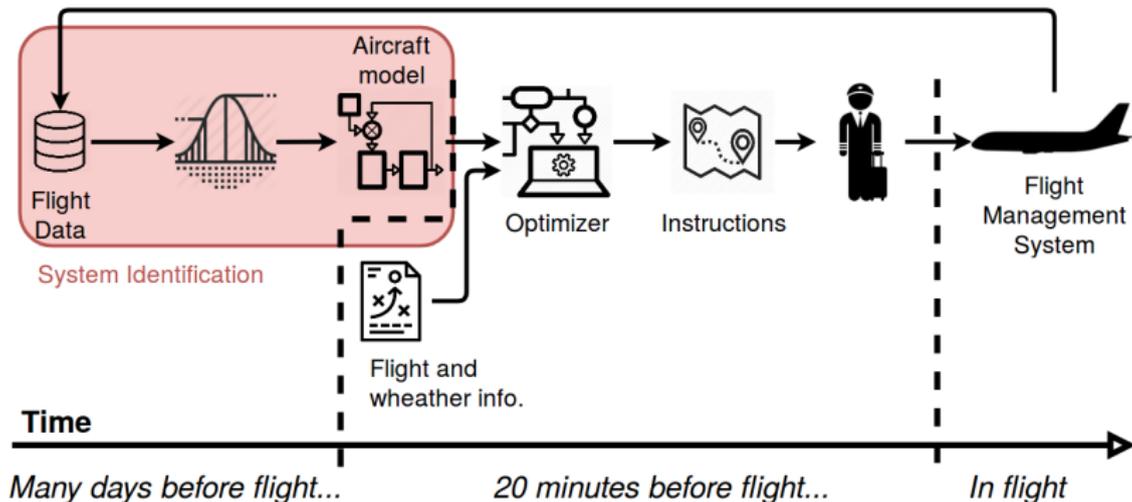


Black box



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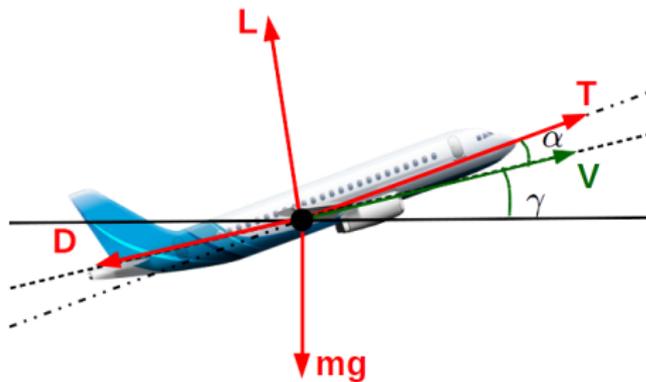
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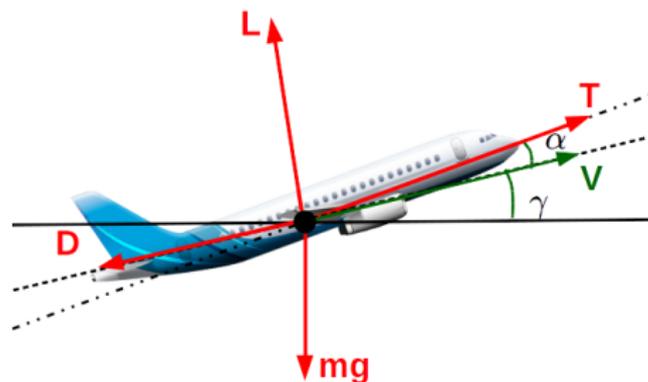
- 1 Context - *Chapter 1*
- 2 System Identification - *Chapter 4*
- 3 Trajectory Acceptability - *Chapters 5 and 6*

SYSTEM IDENTIFICATION

FLIGHT DYNAMICS

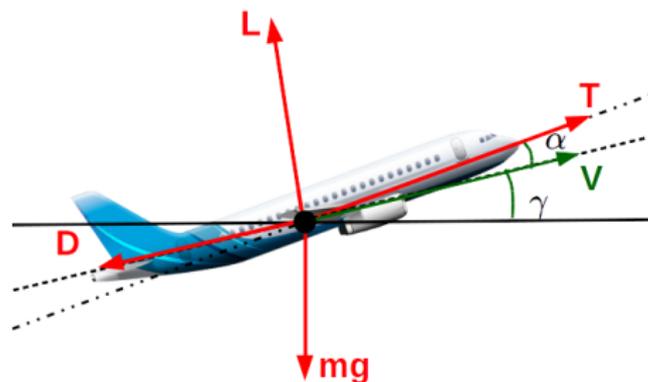


FLIGHT DYNAMICS



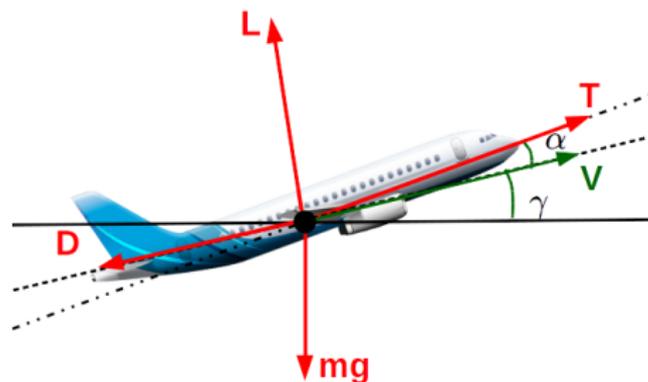
$$\left\{ \begin{array}{l} \dot{h} = V \sin \gamma \\ \dot{V} = \frac{T \cos \alpha - D - mg \sin \gamma}{m} \\ \dot{\gamma} = \frac{T \sin \alpha + L - mg \cos \gamma}{mV} \\ \dot{m} = -\frac{T}{I_{sp}} \end{array} \right.$$

FLIGHT DYNAMICS



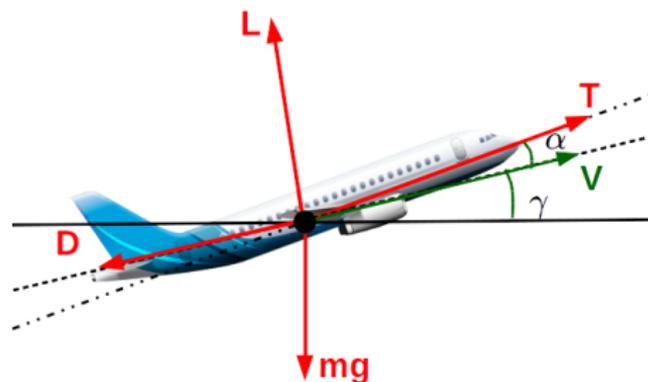
$$\left\{ \begin{array}{l} \dot{h} = V \sin \gamma + \dot{W}_z \\ \dot{V} = \frac{T \cos \alpha - D - mg \sin \gamma - m \dot{W}_{xv}}{m} \\ \dot{\gamma} = \frac{(T \sin \alpha + L) \cos \mu - mg \cos \gamma - m \dot{W}_{zv}}{mV} \\ \dot{m} = -\frac{T}{I_{sp}} \end{array} \right.$$

FLIGHT DYNAMICS



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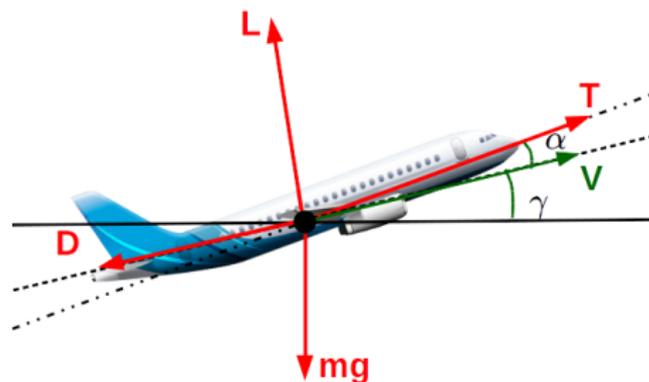
FLIGHT DYNAMICS



States: $x = (h, V, \gamma, m)$

$$\left\{ \begin{array}{l} \dot{h} = V \sin \gamma \\ \dot{V} = \frac{T \cos \alpha - D - mg \sin \gamma}{m} \\ \dot{\gamma} = \frac{T \sin \alpha + L - mg \cos \gamma}{mV} \\ \dot{m} = -\frac{T}{I_{sp}} \end{array} \right.$$

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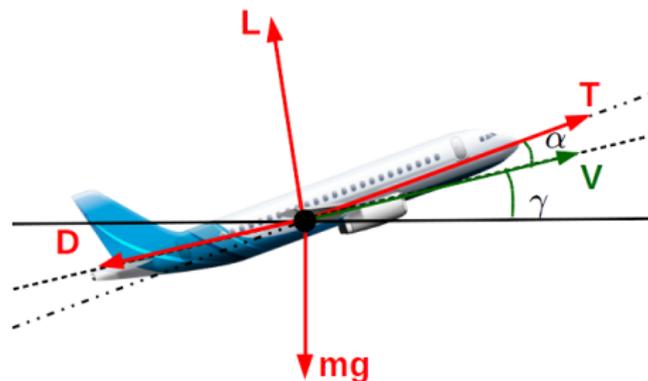


States: $x = (h, V, \gamma, m)$

Controls: $u = (\alpha, N_1)$

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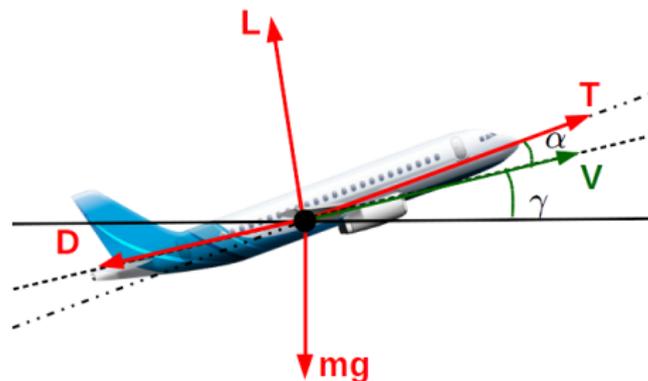
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Unknown functions of x, u

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FLIGHT DYNAMICS



States: $\mathbf{x} = (h, V, \gamma, m)$

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Unknown functions of \mathbf{x}, \mathbf{u}

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PHYSICAL MODELS OF NESTED FUNCTIONS

$$\left\{ \begin{array}{l} T \text{ function of } (N_1, M, \rho) \\ D \text{ function of } (q, M, \alpha) \\ L \text{ function of } (q, M, \alpha) \\ I_{sp} \text{ function of } (SAT, M, h) \end{array} \right.$$

PHYSICAL MODELS OF NESTED FUNCTIONS

$$\left\{ \begin{array}{l} T \text{ function of } (N_1, M, \rho) = \varphi_T(\mathbf{x}, \mathbf{u}) \\ D \text{ function of } (q, M, \alpha) = \varphi_D(\mathbf{x}, \mathbf{u}) \\ L \text{ function of } (q, M, \alpha) = \varphi_L(\mathbf{x}, \mathbf{u}) \\ I_{sp} \text{ function of } (SAT, M, h) = \varphi_{I_{sp}}(\mathbf{x}, \mathbf{u}) \end{array} \right.$$

PHYSICAL MODELS OF NESTED FUNCTIONS

$$\left\{ \begin{array}{l} T(\mathbf{x}, \mathbf{u}, \quad) = N_1 \times P_T(\rho, M) \\ D(\mathbf{x}, \mathbf{u}, \quad) = q \times P_D(\alpha, M) \\ L(\mathbf{x}, \mathbf{u}, \quad) = q \times P_L(\alpha, M) \\ I_{sp}(\mathbf{x}, \mathbf{u}, \quad) = SAT \times P_{Isp}(h, M) \end{array} \right.$$

PHYSICAL MODELS OF NESTED FUNCTIONS

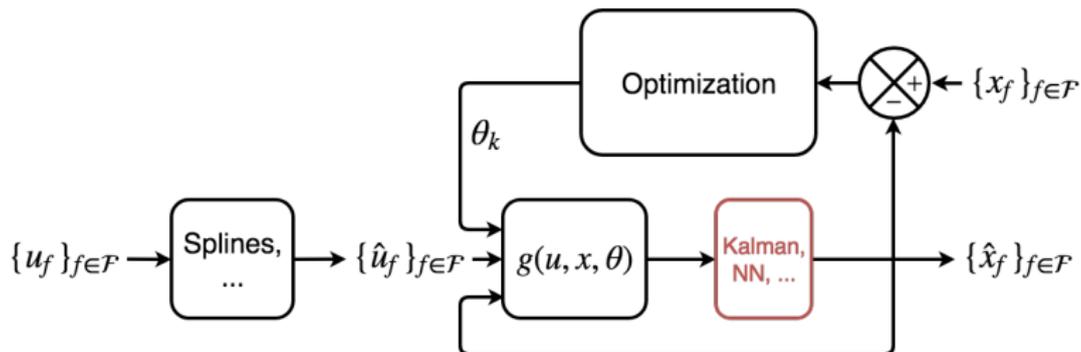
$$\left\{ \begin{array}{l} T(\mathbf{x}, \mathbf{u}, \boldsymbol{\theta}_T) = N_1 \times P_T(\rho, M) = X_T \cdot \boldsymbol{\theta}_T \\ D(\mathbf{x}, \mathbf{u}, \boldsymbol{\theta}_D) = q \times P_D(\alpha, M) = X_D \cdot \boldsymbol{\theta}_D \\ L(\mathbf{x}, \mathbf{u}, \boldsymbol{\theta}_L) = q \times P_L(\alpha, M) = X_L \cdot \boldsymbol{\theta}_L \\ I_{sp}(\mathbf{x}, \mathbf{u}, \boldsymbol{\theta}_{Isp}) = SAT \times P_{Isp}(h, M) = X_{Isp} \cdot \boldsymbol{\theta}_{Isp} \end{array} \right.$$

$$X_T = N_1 \begin{pmatrix} 1 \\ \rho \\ M \\ \rho^2 \\ \rho M \\ M^2 \\ \vdots \end{pmatrix}, X_D = X_L = q \begin{pmatrix} 1 \\ \alpha \\ M \\ \alpha^2 \\ \alpha M \\ M^2 \\ \vdots \end{pmatrix}, X_{Isp} = SAT \begin{pmatrix} 1 \\ h \\ M \\ h^2 \\ hM \\ M^2 \\ \vdots \end{pmatrix}.$$

STATE-OF-THE-ART - [JATEGAONKAR, 2006]

- Output-Error Method
- Filter-Error Method

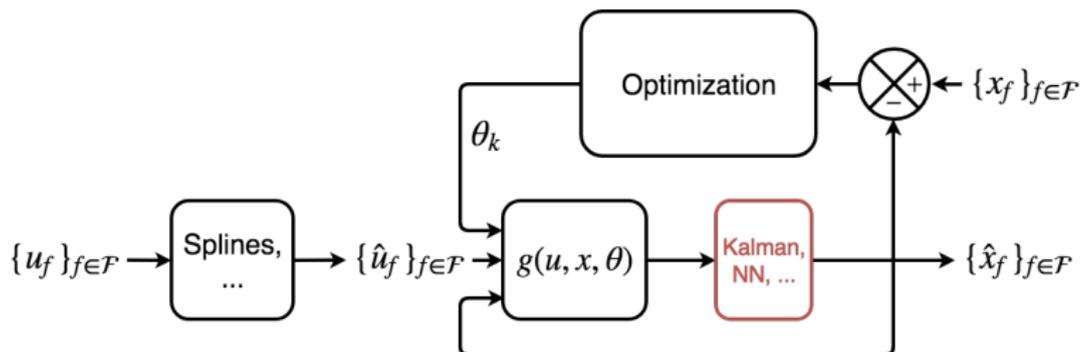
} Less scalable to many trajectories



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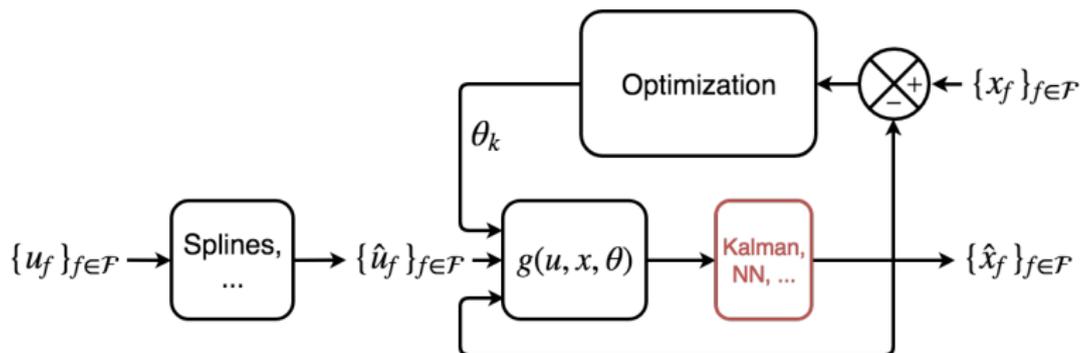
- **Equation-Error Method**

$$\dot{\mathbf{x}}(t) = \mathbf{g}(\mathbf{u}(t), \mathbf{x}(t), \boldsymbol{\theta}) + \boldsymbol{\varepsilon}(t), \quad t \in [0, t_f]$$

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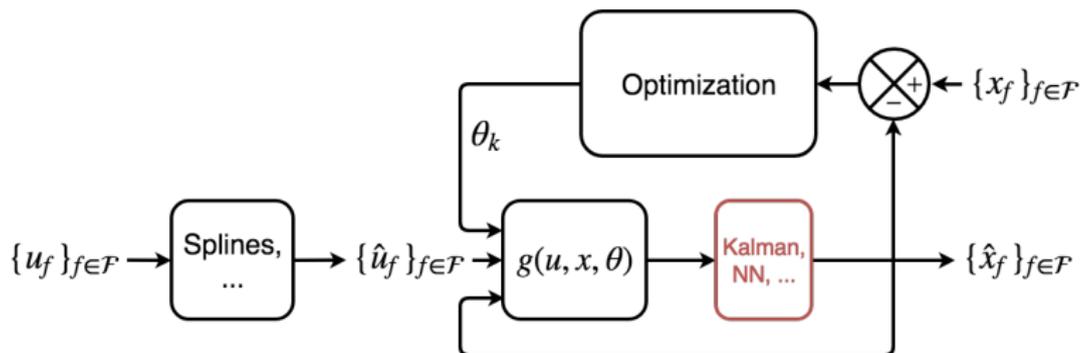
- **Equation-Error Method**

$$\dot{x}_i = g(u_i, x_i, \theta) + \varepsilon_i, \quad i = 1, \dots, N$$

STATE-OF-THE-ART - [JATEGAONKAR, 2006]

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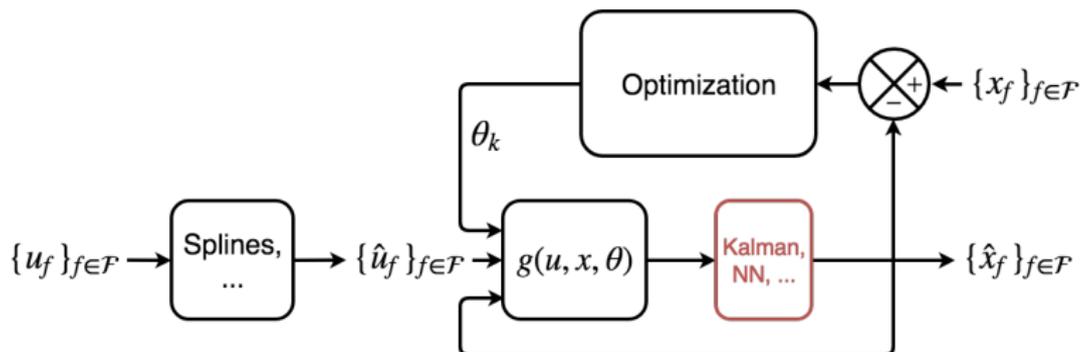
- **Equation-Error Method**

$$\min_{\theta} \sum_{i=1}^N \ell(\dot{x}_i, g(\mathbf{u}_i, \mathbf{x}_i, \theta))$$

STATE-OF-THE-ART - [JATEGAONKAR, 2006]

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- **Equation-Error Method**

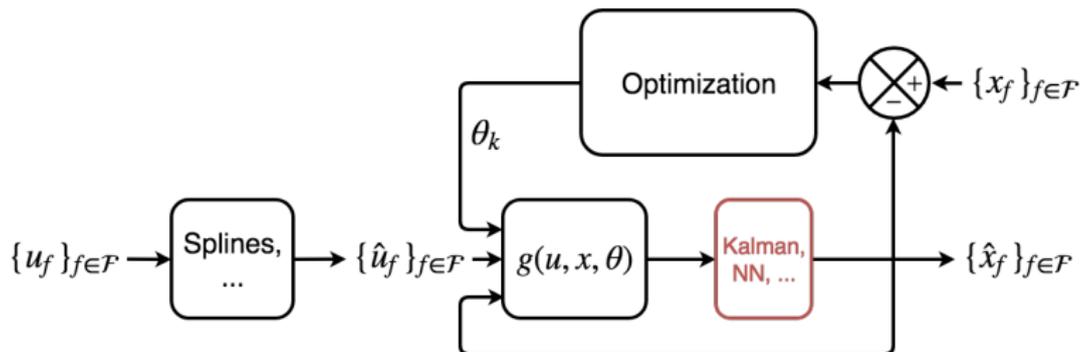
Ex: (Nonlinear) Least-Squares

$$\min_{\theta} \sum_{i=1}^N \left\| \dot{x}_i - g(\mathbf{u}_i, \mathbf{x}_i, \theta) \right\|_2^2$$

STATE-OF-THE-ART - [JATEGAONKAR, 2006]

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} Less scalable to many trajectories



- **Equation-Error Method**

Ex: (Nonlinear) Least-Squares

$$\min_{\theta} \sum_{i=1}^N \left\| Y(\mathbf{u}_i, \mathbf{x}_i, \dot{\mathbf{x}}_i) - G(\mathbf{u}_i, \mathbf{x}_i, \dot{\mathbf{x}}_i, \theta) \right\|_2^2$$

LEVERAGING THE DYNAMICS STRUCTURE

$$\left\{ \begin{array}{l} \dot{h} = V \sin \gamma \\ \dot{V} = \frac{T(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_T) \cos \alpha - D(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_D) - mg \sin \gamma}{m} \\ \dot{\gamma} = \frac{T(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_T) \sin \alpha + L(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_L) - mg \cos \gamma}{mV} \\ \dot{m} = -\frac{T(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_T)}{I_{sp}(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_{Isp})} \end{array} \right.$$

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- Nonlinear in states and controls

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- Nonlinear in states and controls
- Nonlinear in parameters

LEVERAGING THE DYNAMICS STRUCTURE

$$\left\{ \begin{array}{l} \dot{h} = V \sin \gamma \\ m\dot{V} + mg \sin \gamma = T(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_T) \cos \alpha - D(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_D) \\ mV\dot{\gamma} + mg \cos \gamma = T(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_T) \sin \alpha + L(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_L) \\ 0 = T(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_T) + \dot{m}l_{sp}(\mathbf{u}, \mathbf{x}, \boldsymbol{\theta}_{lsp}) \end{array} \right.$$

- Nonlinear in states and controls
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LEVERAGING THE DYNAMICS STRUCTURE

$$\left\{ \begin{array}{l} \dot{h} = V \sin \gamma \\ Y_1 = X_{T1} \cdot \theta_T - X_D \cdot \theta_D + \varepsilon_1 \\ Y_2 = X_{T2} \cdot \theta_T + X_L \cdot \theta_L + \varepsilon_2 \\ Y_3 = X_T \cdot \theta_T + X_{Ispm} \cdot \theta_{Ispm} + \varepsilon_3 \end{array} \right.$$

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MULTI-TASK REGRESSION

Aircraft:

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General:

$$\begin{cases} Y_1 = X_{c,1} \cdot \theta_c + X_1 \cdot \theta_1 + \varepsilon_1 \\ Y_2 = X_{c,2} \cdot \theta_c + X_2 \cdot \theta_2 + \varepsilon_2 \\ \vdots \\ Y_K = X_{c,K} \cdot \theta_c + X_K \cdot \theta_K + \varepsilon_K \end{cases}$$

Coupling parameters , **Task specific parameters**

Multi-task Linear Least-Squares:

$$\min_{\theta} \sum_{k=1}^K \sum_{i=1}^N (Y_{k,i} - X_{c,k,i} \cdot \theta_c - X_{k,i} \cdot \theta_k)^2$$

MULTI-TASK REGRESSION

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Coupling parameters , **Task specific parameters**

Multi-task Linear Least-Squares:

Block-sparse Coupling Structure

$$\min_{\theta} \sum_{i=1}^N \left\| \begin{pmatrix} Y_{1,i} \\ \vdots \\ Y_{K,i} \end{pmatrix} - \begin{pmatrix} X_{c,1,i}^\top & X_{1,i}^\top & 0 & 0 & \dots & 0 \\ X_{c,2,i}^\top & 0 & X_{2,i}^\top & 0 & \dots & 0 \\ \vdots & 0 & 0 & \ddots & 0 & 0 \\ X_{c,K,i}^\top & 0 & 0 & \dots & 0 & X_{K,i}^\top \end{pmatrix} \begin{pmatrix} \theta_c \\ \theta_1 \\ \vdots \\ \theta_K \end{pmatrix} \right\|_2^2$$

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$$\min_{\theta} \sum_{i=1}^N \|Y_i - X_i \theta\|_2^2$$

with $\theta = (\theta_c, \theta_1, \dots, \theta_K) \in \mathbb{R}^p$, $p = d_c + \sum_{k=1}^K d_k$,
 $Y_i \in \mathbb{R}^K$ and $X_i \in \mathbb{R}^{K \times p}$.

FEATURE SELECTION

Our model:

$$T = N_1(\theta_{T,1} + \theta_{T,2}\rho + \theta_{T,3}M + \theta_{T,4}\rho^2 + \theta_{T,5}\rho M + \theta_{T,6}M^2 + \theta_{T,7}\rho^3 + \theta_{T,8}\rho^2 M + \theta_{T,9}\rho M^2 + \theta_{T,10}M^3 + \theta_{T,11}\rho^4 + \theta_{T,12}\rho^3 M + \theta_{T,13}\rho^2 M^2 + \theta_{T,14}\rho M^3 + \theta_{T,15}M^4).$$

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⇒ High risk of overfitting

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Sparse models are:

- Less susceptible to overfitting,
- More compliant with physical models,
- More interpretable,
- Lighter/Faster.

BLOCK-SPARSE LASSO

Lasso [Tibshirani, 1994]: $\{(X_i, Y_i)\}_{i=1}^N \subset \mathbb{R}^{d+1}$ i.i.d sample,

$$\min_{\theta} \sum_{i=1}^N (Y_i - X_i \cdot \theta)^2 + \lambda \|\theta\|_1.$$

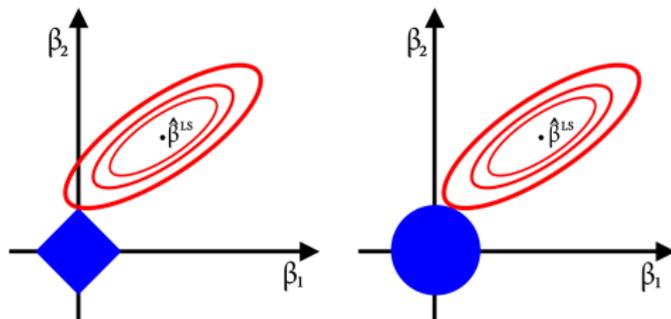


FIGURE: ¹Sparsity induced by L^1 norm in Lasso.

BLOCK-SPARSE LASSO

$$\min_{\theta} \sum_{k=1}^K \sum_{i=1}^N (Y_{k,i} - X_{c,k,i} \cdot \theta_c - X_{k,i} \cdot \theta_k)^2 + \lambda_c \|\theta_c\|_1 + \sum_{k=1}^K \lambda_k \|\theta_k\|_1$$

BLOCK-SPARSE LASSO

Block-sparse structure preserved \rightsquigarrow **Equivalent to Lasso problem**

$$\min_{\theta} \sum_{k=1}^K \sum_{i=1}^N (Y_{k,i} - X_{c,k,i} \cdot \theta_c - X_{k,i} \cdot \theta_k)^2 + \lambda_c \|\theta_c\|_1 + \sum_{k=1}^K \lambda_k \|\theta_k\|_1$$

BLOCK-SPARSE LASSO

Block-sparse structure preserved \rightsquigarrow **Equivalent to Lasso problem**

$$\min_{\beta} \sum_{i=1}^N \|Y_i - B_i \beta\|_2^2 + \lambda_c \|\beta\|_1$$

with $\beta = (\boldsymbol{\theta}_c, \frac{\lambda_1}{\lambda_c} \boldsymbol{\theta}_1, \dots, \frac{\lambda_K}{\lambda_c} \boldsymbol{\theta}_K) \in \mathbb{R}^p$, $p = d_c + \sum_{k=1}^K d_k$,
 $Y_i \in \mathbb{R}^K$ and $B_i \in \mathbb{R}^{K \times p}$.

BLOCK-SPARSE LASSO

Block-sparse structure preserved \rightsquigarrow **Equivalent to Lasso problem**

$$\min_{\boldsymbol{\theta}} \sum_{i=1}^N \|Y_i - X_i \boldsymbol{\theta}\|_2^2 + \lambda_c \|\boldsymbol{\theta}\|_1$$

with $\boldsymbol{\theta} = (\boldsymbol{\theta}_c, \boldsymbol{\theta}_1, \dots, \boldsymbol{\theta}_K) \in \mathbb{R}^p$, $p = d_c + \sum_{k=1}^K d_k$,

$Y_i \in \mathbb{R}^K$ and $X_i \in \mathbb{R}^{K \times p}$,

In practice, we choose $\lambda_k = \lambda_c$, for all $k = 1, \dots, 3$ and

$$X_i = \begin{pmatrix} X_{T1,i}^\top & -X_{D,i}^\top & 0 & 0 \\ X_{T2,i}^\top & 0 & X_{L,i}^\top & 0 \\ X_{T,i}^\top & 0 & 0 & X_{lspm,i}^\top \end{pmatrix}, \quad Y_i = \begin{pmatrix} Y_{1,i} \\ Y_{2,i} \\ Y_{3,i} \end{pmatrix}$$

BOOTSTRAP IMPLEMENTATION

High correlations between features...

BOOTSTRAP IMPLEMENTATION

High correlations between features...

⇒ Inconsistent selections via the lasso !

PROBLEM WITH INTRA-GROUP CORRELATIONS

$$\min_{\boldsymbol{\theta}} \sum_{i=1}^N \|Y_i - X_i \boldsymbol{\theta}\|_2^2 + \lambda_c \|\boldsymbol{\theta}\|_1 \Rightarrow \hat{\boldsymbol{\theta}}_T = \hat{\boldsymbol{\theta}}_{lsp} = \mathbf{0}!$$

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PROBLEM WITH INTRA-GROUP CORRELATIONS

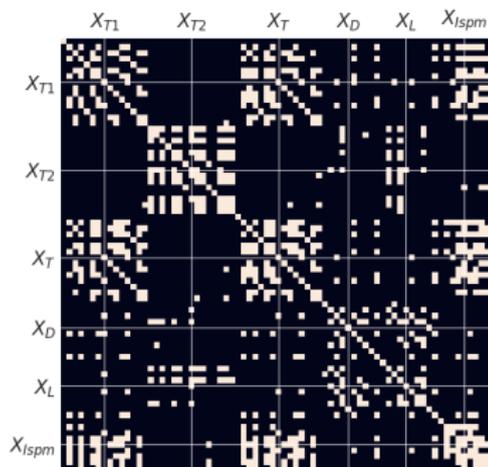
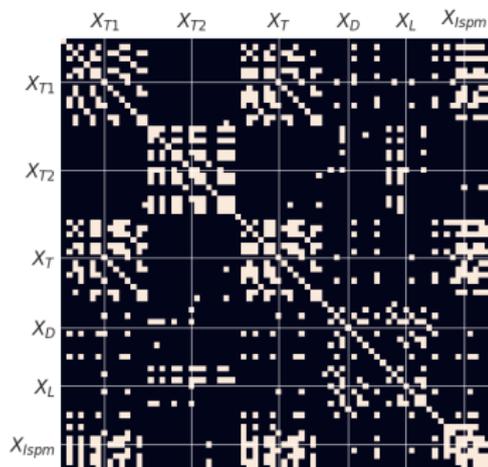


FIGURE: Features correlations higher than 0.9 in absolute value in white.

PROBLEM WITH INTRA-GROUP CORRELATIONS



$\Rightarrow \theta \mapsto \sum_{i=1}^N \|Y_i - X_i\theta\|_2^2$ not injective...

Ill-posed problem !

FIGURE: Features correlations higher than 0.9 in absolute value in white.

PROBLEM WITH INTRA-GROUP CORRELATIONS

$$\left\{ \begin{array}{l} Y_1 = X_{T1} \cdot \boldsymbol{\theta}_T - X_D \cdot \boldsymbol{\theta}_D + \varepsilon_1 \\ Y_2 = X_{T2} \cdot \boldsymbol{\theta}_T + X_L \cdot \boldsymbol{\theta}_L + \varepsilon_2 \\ 0 = X_T \cdot \boldsymbol{\theta}_T + X_{Ispm} \cdot \boldsymbol{\theta}_{Isp} + \varepsilon_3 \end{array} \right.$$

$$\min_{\boldsymbol{\theta}} \sum_{i=1}^N \|Y_i - X_i \boldsymbol{\theta}\|_2^2 + \lambda_c \|\boldsymbol{\theta}\|_1$$

Prior model \tilde{I}_{sp} from Roux [2005] $\rightsquigarrow \tilde{I}_{sp,i} = \tilde{I}_{sp}(\mathbf{u}_i, \mathbf{x}_i)$, $i = 1, \dots, N$.

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$$\min_{\boldsymbol{\theta}} \sum_{i=1}^N \left(\|Y_i - X_i \boldsymbol{\theta}\|_2^2 + \lambda_t \|\tilde{I}_{sp,i} - X_{Isp,i} \cdot \boldsymbol{\theta}_{Isp}\|_2^2 \right) + \lambda_c \|\boldsymbol{\theta}\|_1$$

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PROBLEM WITH INTRA-GROUP CORRELATIONS

$$\begin{cases} Y_1 &= X_{T1} \cdot \boldsymbol{\theta}_T & - X_D \cdot \boldsymbol{\theta}_D & + \varepsilon_1 \\ Y_2 &= X_{T2} \cdot \boldsymbol{\theta}_T & + X_L \cdot \boldsymbol{\theta}_L & + \varepsilon_2 \\ 0 &= X_T \cdot \boldsymbol{\theta}_T & + X_{Ispm} \cdot \boldsymbol{\theta}_{Isp} & + \varepsilon_3 \\ \sqrt{\lambda_t} \tilde{I}_{sp} &= \sqrt{\lambda_t} X_{Isp} \cdot \boldsymbol{\theta}_{Isp} & & + \varepsilon_4 \end{cases}$$

$$\min_{\boldsymbol{\theta}} \sum_{i=1}^N \|\tilde{Y}_i - \tilde{X}_i \boldsymbol{\theta}\|_2^2 + \lambda_c \|\boldsymbol{\theta}\|_1$$

$$\tilde{Y}_i = \begin{pmatrix} Y_{1,i} \\ Y_{2,i} \\ 0 \\ \sqrt{\lambda_t} \tilde{I}_{sp,i} \end{pmatrix}, \quad \tilde{X}_i = \begin{pmatrix} X_{T1,i}^\top & -X_{D,i}^\top & 0 & 0 \\ X_{T2,i}^\top & 0 & X_{L,i}^\top & 0 \\ X_{T,i}^\top & 0 & 0 & X_{Ispm,i}^\top \\ 0 & 0 & 0 & \sqrt{\lambda_t} X_{Isp,i}^\top \end{pmatrix},$$

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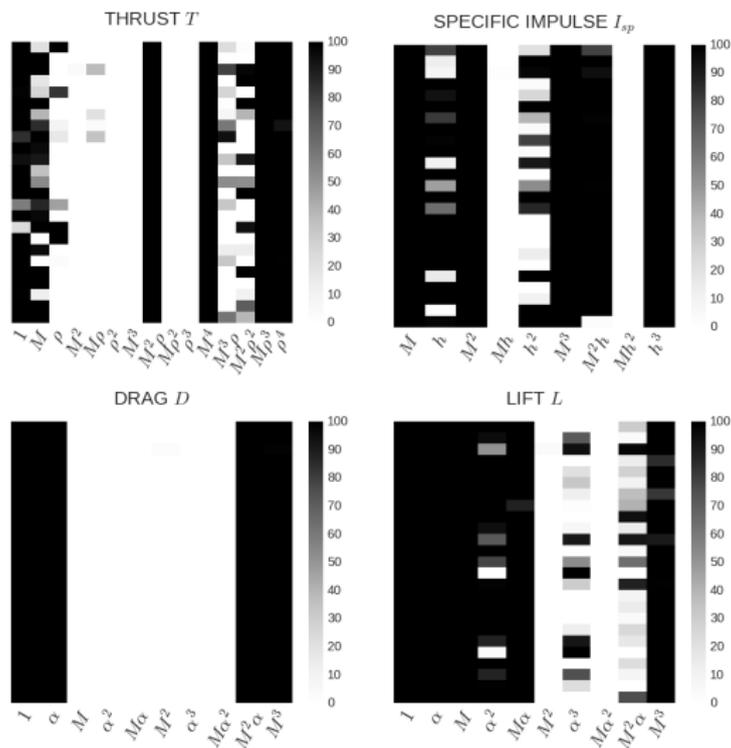
FEATURE SELECTION RESULTS

- 25 different B737-800,
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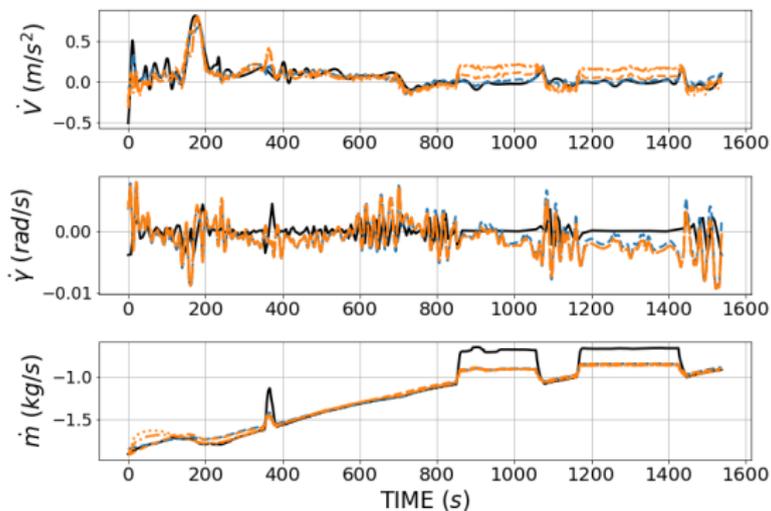
- 25 different B737-800,
- 10 471 flights = 8 261 619 observations,
- Block sparse Bolasso used for T , D , L and I_{sp} ,
- We expect similar model structures,

FEATURE SELECTION RESULTS

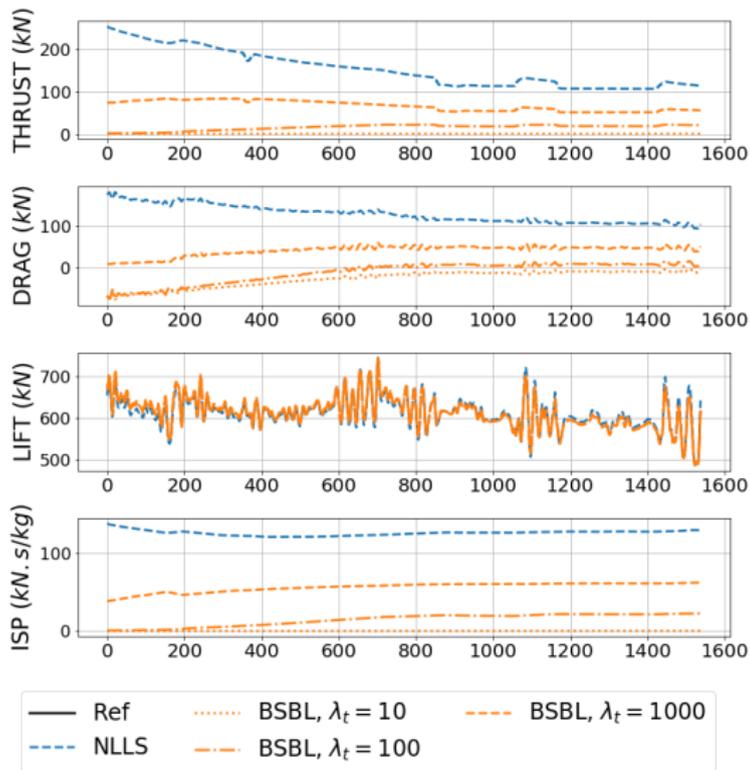


Feature selection results for the thrust, drag, lift and specific impulse models.

ACCURACY OF DYNAMICS PREDICTIONS



REALISM OF HIDDEN ELEMENTS



FLIGHT RESIMULATION

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Another possible dynamic criterion:

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where $\|\cdot\|_{\mathbf{u}}$, $\|\cdot\|_{\mathbf{x}}$ denote scaling norms.

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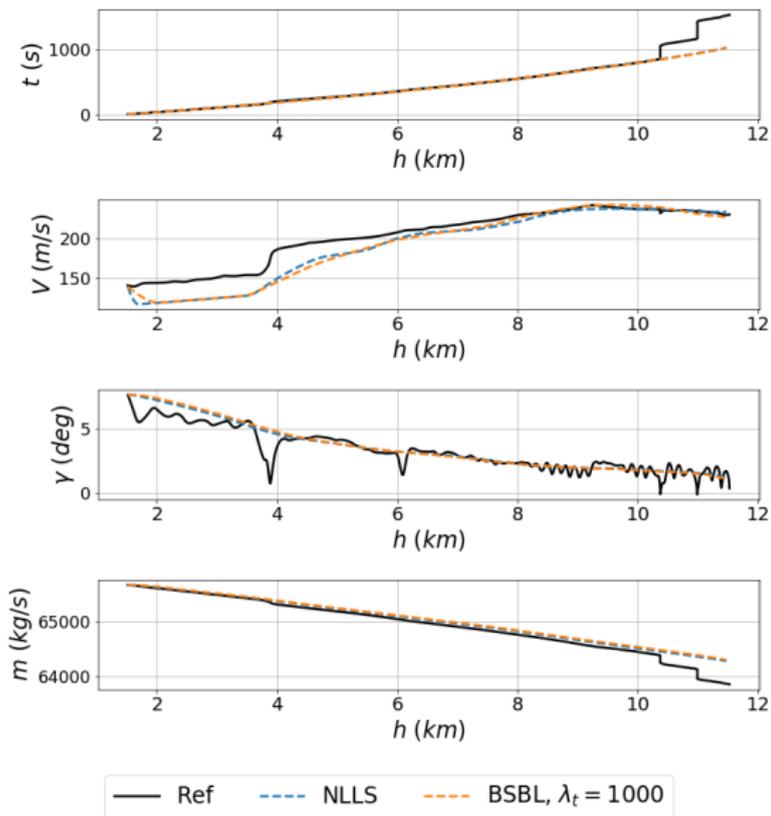
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For practical applications: $t \leftrightarrow h$

FLIGHT RESIMULATION



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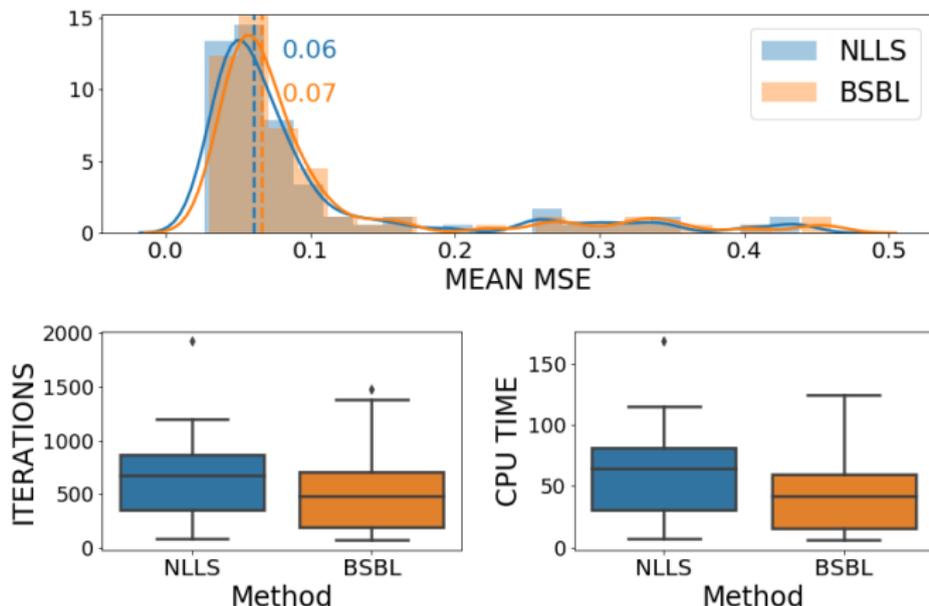


FIGURE: Distribution of the off-sample simulation error and boxplot of the optimization number of iterations and CPU time.

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 - Faster convergence when applied to control problems.

TRAJECTORY ACCEPTABILITY

$$\begin{aligned} & \min_{(\mathbf{x}, \mathbf{u}) \in \mathbb{X} \times \mathbb{U}} \int_0^{t_f} C(\mathbf{u}(t), \mathbf{x}(t)) dt, \\ \text{s.t. } & \begin{cases} \dot{\mathbf{x}}(t) = \hat{\mathbf{g}}(\mathbf{u}(t), \mathbf{x}(t)), & \text{a.e. } t \in [0, t_f], \\ \text{Other constraints...} \end{cases} \end{aligned} \quad (\text{AOCP})$$

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$\Rightarrow \hat{\mathbf{z}} = (\hat{\mathbf{x}}, \hat{\mathbf{u}})$ solution of (AOCP).

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Pilots acceptance



Air Traffic Control²

How can we quantify the closeness from the optimized trajectory to the set of real flights?

OPTIMIZED TRAJECTORY LIKELIHOOD

Assumption: We suppose that the real flights are observations of the same functional random variable $Z = (Z_t)$ valued in $\mathcal{C}(\mathbb{T}, E)$, with E compact subset of \mathbb{R}^d and $\mathbb{T} = [0, t_f]$.

How likely is it to draw the optimized trajectory from the law of Z ?

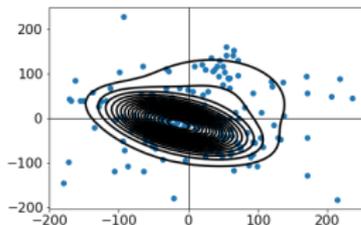
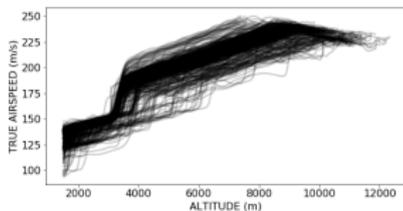
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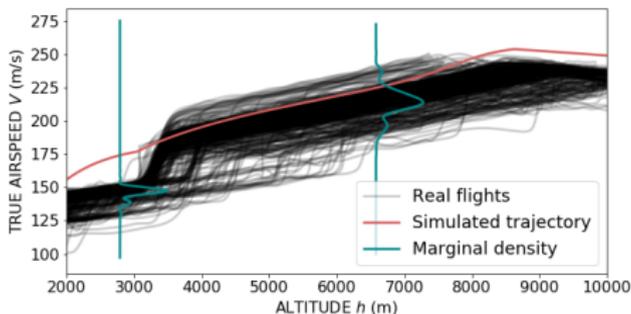
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HOW TO APPLY THIS TO FUNCTIONAL DATA?

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- Standard approach in Functional Data Analysis: use Functional Principal Component Analysis to decompose the data in a small number of coefficients
- **Or: we can use the marginal densities**



HOW DO WE AGGREGATE THE MARGINAL LIKELIHOODS?

- f_t marginal density of Z , i.e. probability density function of Z_t ,
- \mathbf{y} new trajectory,
- $f_t(\mathbf{y}(t))$ marginal likelihood of \mathbf{y} at t , i.e. likelihood of observing $Z_t = \mathbf{y}(t)$.

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MEAN MARGINAL LIKELIHOOD

$$\text{MML}(Z, \mathbf{y}) = \frac{1}{t_f} \int_0^{t_f} \psi[f_t, \mathbf{y}(t)] dt,$$

where $\psi : L^1(E, \mathbb{R}_+) \times \mathbb{R} \rightarrow [0; 1]$ is a continuous scaling map,

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Possible scalings are the normalized density

$$\psi[f_t, \mathbf{y}(t)] := \frac{f_t(\mathbf{y}(t))}{\max_{z \in E} f_t(z)},$$

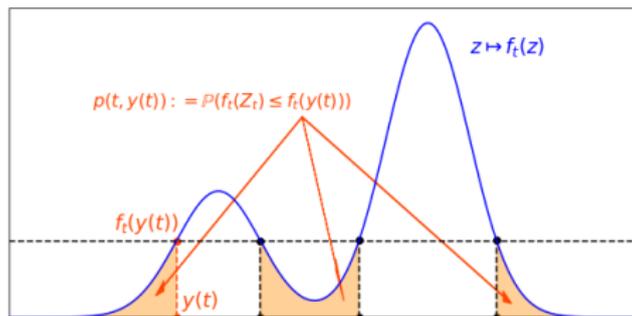
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or the confidence level

$$\psi[f_t, \mathbf{y}(t)] := \mathbb{P}(f_t(Z_t) \leq f_t(\mathbf{y}(t))).$$



HOW DO WE DEAL WITH SAMPLED CURVES?

In practice, the m trajectories are sampled at variable discrete times:

$$\mathcal{T}^D := \left\{ (t_j^r, z_j^r) \right\}_{\substack{1 \leq j \leq n \\ 1 \leq r \leq m}} \subset \mathbb{T} \times E, \quad z_j^r := \mathbf{z}^r(t_j^r),$$
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Hence, we approximate the MML using a Riemann sum which aggregates consistent estimators $\hat{f}_{\tilde{t}_j}^m$ of the marginal densities $f_{\tilde{t}_j}$:

$$\text{EMML}_m(\mathcal{T}^D, \mathcal{Y}) := \frac{1}{t_f} \sum_{j=1}^{\tilde{n}} \psi[\hat{f}_{\tilde{t}_j}^m, y_j] \Delta \tilde{t}_j.$$

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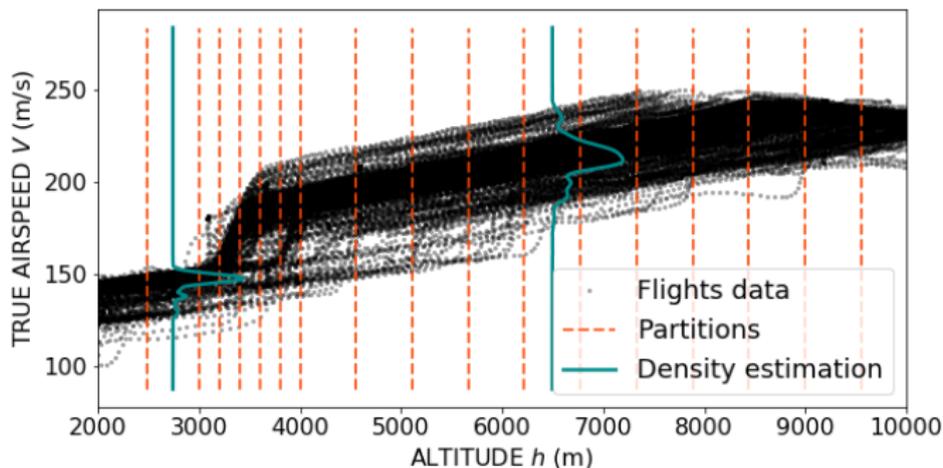
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- 1 We can apply SOA conditional density estimation techniques, such as LS-CDE [Sugiyama et al., 2010],
 - 2 **We can use a fine partitioning of the time domain.**

PARTITION BASED MARGINAL DENSITY ESTIMATION



Idea: to average in time the marginal densities over small bins by applying classical multivariate density estimation techniques to each subset.

CONSISTENCY

We denote by:

- $\Theta : \mathcal{S} \rightarrow L^1(E, \mathbb{R}_+)$ multivariate density estimation statistic,
- $\mathcal{S} = \{(z_k)_{k=1}^N \in E^N : N \in \mathbb{N}^*\}$ set of finite sequences,

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- $\hat{f}_t^m := \Theta[\mathcal{T}_t^m]$ estimator trained using \mathcal{T}_t^m .

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ASSUMPTION 1 - POSITIVE TIME DENSITY

$\nu \in L^\infty(E, \mathbb{R}_+)$ density function of T , s.t.

$$\nu_+ := \operatorname{ess\,sup}_{t \in \mathbb{T}} \nu(t) < \infty, \quad \nu_- := \operatorname{ess\,inf}_{t \in \mathbb{T}} \nu(t) > 0.$$

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Function $(t, z) \in \mathbb{T} \times E \mapsto f_t(z)$ is continuous and

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ASSUMPTION 3 - SHRINKING BINS

The homogeneous partition $\{B_\ell^m\}_{\ell=1}^{q_m}$ of $[0; t_f]$, with binsize b_m , is s.t.

$$\lim_{m \rightarrow \infty} b_m = 0, \quad \lim_{m \rightarrow \infty} m b_m = \infty.$$

CONSISTENCY

ASSUMPTION 4 - I.I.D. CONSISTENCY

- \mathcal{G} arbitrary family of probability density functions on E , $\rho \in \mathcal{G}$,
- S_ρ^N **i.i.d** sample of **size N** drawn from ρ valued in \mathcal{S} .

The estimator obtained by applying Θ to S_ρ^N , denoted by

$$\hat{\rho}^N := \Theta[S_\rho^N] \in L^1(E, \mathbb{R}_+),$$

is a (pointwise) consistent density estimator, uniformly in ρ :

For all $z \in E$, $\varepsilon > 0$, $\alpha_1 > 0$, there is $N_{\varepsilon, \alpha_1} > 0$ such that, for any $\rho \in \mathcal{G}$,

$$N \geq N_{\varepsilon, \alpha_1} \Rightarrow \mathbb{P} \left(\left| \hat{\rho}^N(z) - \rho(z) \right| < \varepsilon \right) > 1 - \alpha_1.$$

CONSISTENCY

THEOREM 1

Under assumptions 1 to 4, for any $z \in E$ and $t \in \mathbb{T}$, $\hat{f}_{\ell^m(t)}^m(z)$ consistently approximates the marginal density $f_t(z)$ as the number of curves m grows:

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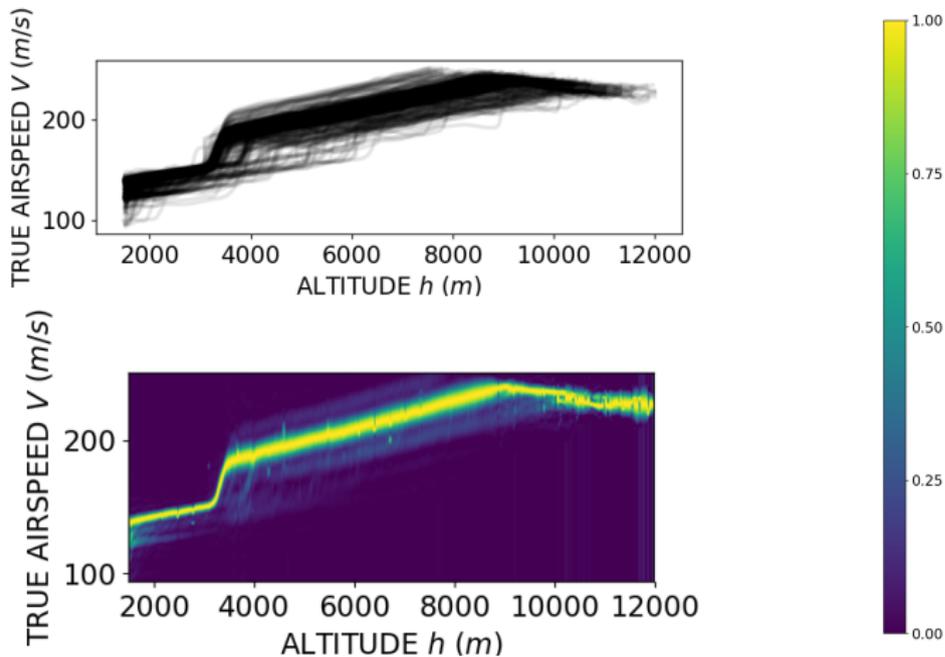
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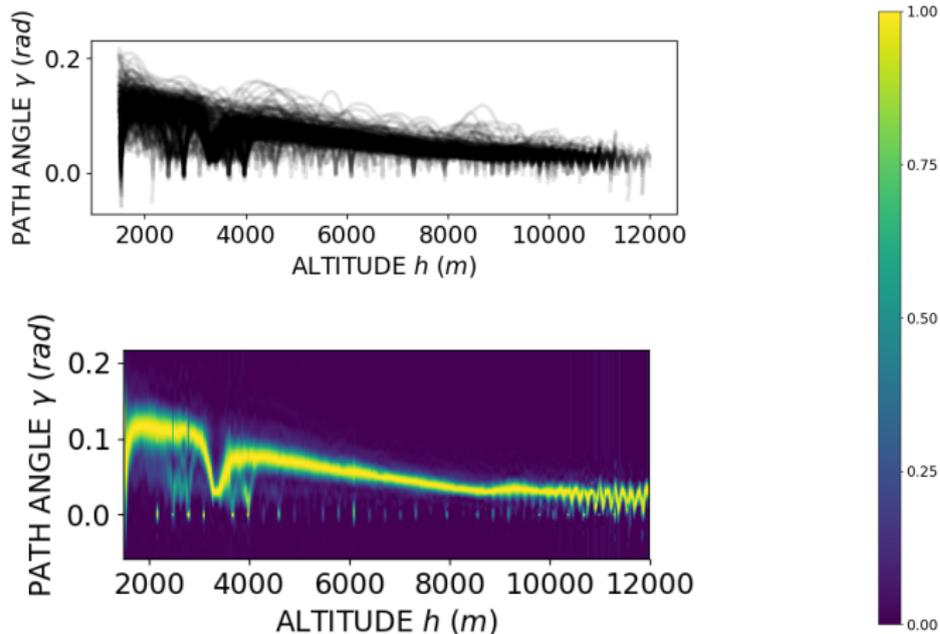
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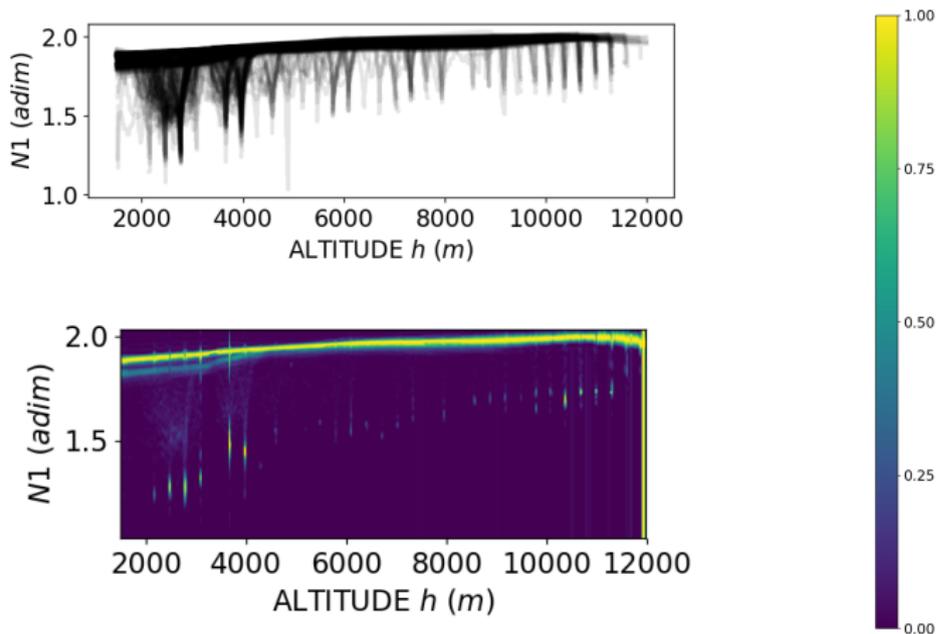
MARGINAL DENSITY ESTIMATION RESULTS



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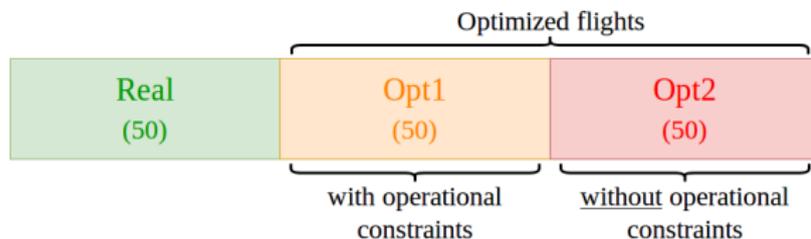
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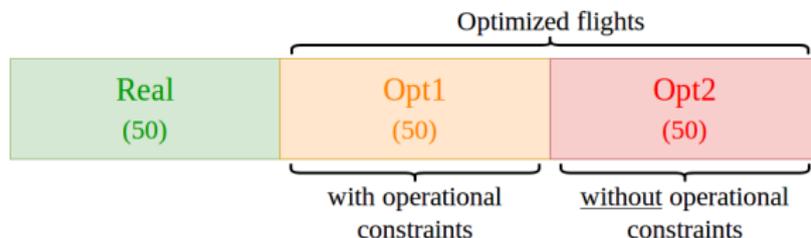
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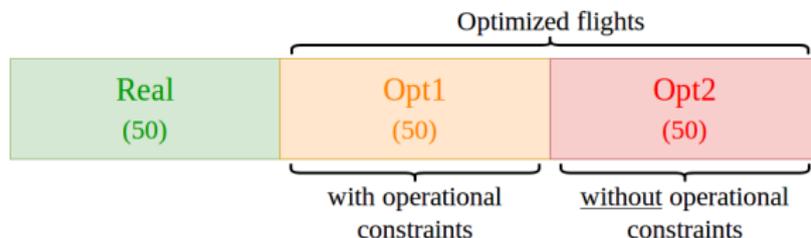


- Discrimination power comparison with (gmm-)FPCA and (integrated) LS-CDE:

VAR.	ESTIMATED LIKELIHOODS		
	REAL	OPT1	OPT2
MML	0.63 ± 0.07	0.43 ± 0.08	0.13 ± 0.02
FPCA	0.16 ± 0.12	$6.4E-03 \pm 3.8E-03$	$3.6E-03 \pm 5.4E-03$
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VAR.	ESTIMATED LIKELIHOODS			TR. TIME
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FPCA	0.16 ± 0.12	$6.4E-03 \pm 3.8E-03$	$3.6E-03 \pm 5.4E-03$	20s
LS-CDE	0.77 ± 0.05	0.68 ± 0.04	0.49 ± 0.06	14H

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The MML can be used not only to assess the optimization solutions, but also to penalize the optimization itself:

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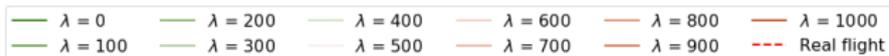
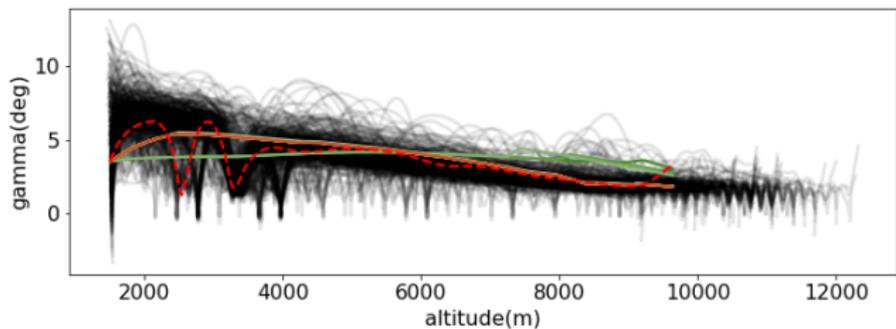
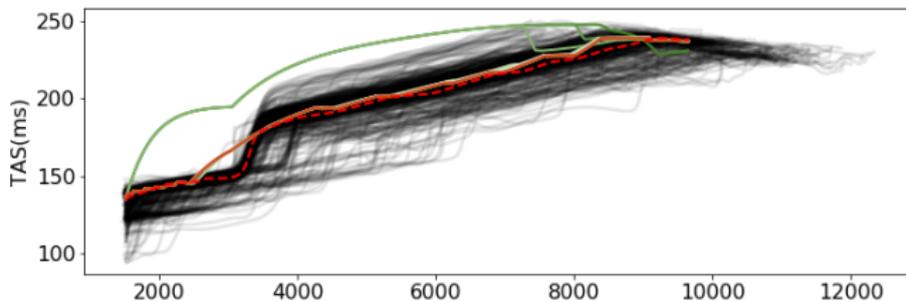
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- λ sets trade-off between a fuel minimization and a likelihood maximization,

PENALTY EFFECT



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- 4 Particular Adaptive Kernel and Gaussian mixture implementation,
 - Showed that it can be used in optimal control problems to obtain solutions close to optimal, and still realistic.

REFERENCES

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ACCURACY OF DYNAMICS PREDICTIONS

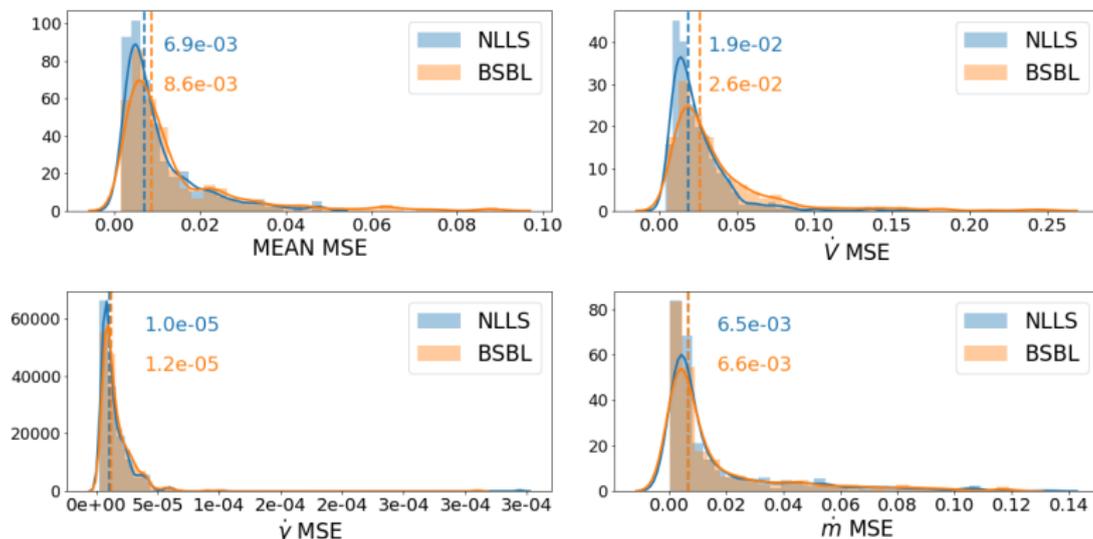


FIGURE: Leave-one-out off-sample errors distributions for nonlinear least-squares NLLS and block-sparse bolasso BSBL. Median errors are annotated and marked by dashed vertical lines.

STRUCTURED FEATURE SELECTION

STATE-OF-THE-ART

Other methods	Difference with Block-sparse Lasso
Group Lasso [Yuan and Lin, 2005]	Groups sparsity is fixed <i>a priori</i> ,
Sparse Group Lasso [Friedman et al., 2010]	Sparsity induced <u>only</u> within group,
Multi-task Lasso [Obozinski et al., 2006]	Not same pattern for every task.

THEOREM (BOLASSO CONSISTENCY - BACH [2008])

For $\lambda = \lambda_0 N^{-\frac{1}{2}}$ and $\lambda_0 > 0$, assume that

- (H1) the cumulant generating functions $\mathbb{E} [\exp(s\|X\|_2^2)]$ and $\mathbb{E} [\exp(s\|Y\|_2^2)]$ are finite for some $s > 0$.
- (H2) the joint matrix of second order moments $Q = \mathbb{E} [XX^\top] \in \mathbb{R}^{p \times p}$ is invertible.
- (H3) $\mathbb{E} [Y|X] = X \cdot \theta$ and $\text{Var} [Y|X] = \sigma^2$ a.s. for some $\theta \in \mathbb{R}^p$ and $\sigma \in \mathbb{R}_+^*$.

Then, for any $b > 0$, the probability that algorithm 1 does not exactly select the correct model has the following upper bound:

$$\mathbb{P} [J \neq J^*] \leq bA_1 e^{-A_2 N} + A_3 \frac{\log N}{N^{1/2}} + A_4 \frac{\log b}{b},$$

where $A_1, A_2, A_3, A_4 > 0$.

GENERALIZED TIKHONOV REGULARIZATION OF ISP

Equivalent to $\|\Gamma(\boldsymbol{\theta} - \tilde{\boldsymbol{\theta}})\|_2^2$ with $\Gamma_i = (\underbrace{0, \dots, 0}_{d_T + d_D + d_L}, X_{isp}^\top)$ and

$$\Gamma_i \tilde{\boldsymbol{\theta}} = \tilde{l}_{sp,i}.$$

MML CONSISTENCY FOR STANDARD KERNEL ESTIMATOR

ASSUMPTION 5

The function $(t, z) \in \mathbb{T} \times E \mapsto f_t(z)$ is $\mathcal{C}^4(E)$ in z and $\mathcal{C}^1(\mathbb{T})$ in t ; the Lipschitz constant of the function

$$t \mapsto \frac{d^2 f_t}{dz^2}(z) := f_t''(z)$$

is denoted by $L'' > 0$: for any $z \in E$ and $t_1, t_2 \in \mathbb{T}$,

$$|f_{t_1}''(z) - f_{t_2}''(z)| \leq L'' |t_1 - t_2|.$$

MML CONSISTENCY FOR STANDARD KERNEL ESTIMATOR

$$\sigma_{K_\sigma}^2 = \int w^2 K_\sigma(w) dw = \sigma^2 \int w^2 K(w) dw = \sigma^2 \sigma_K^2,$$

$$\sigma_{K_\sigma^2}^2 = \int w^2 K_\sigma(w)^2 dw = \sigma \int w^2 K(w)^2 dw = \sigma \sigma_{K^2}^2,$$

$$R(K_\sigma) = \int K_\sigma(w)^2 dw = \frac{1}{\sigma} \int K(w)^2 dw = \frac{1}{\sigma} R(K).$$

THEOREM 2

Under assumptions 1, 3 and 5, if $\hat{f}_{\ell^m(t)}^m$ is a KDE where the kernel K and the bandwidth $\sigma := \sigma_m$ are deterministic, such that $\sigma_K < \infty$, $\sigma_{K^2} < \infty$, $R(K) < \infty$ and if

$$\lim_{m \rightarrow \infty} \sigma_m = 0, \quad \lim_{m \rightarrow \infty} mb_m \sigma_m = +\infty,$$

then

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[(\hat{f}_{\ell^m(t)}^m(z) - f_t(z))^2 \right] = 0.$$

THEOREM 1 PROOF SKETCH

$$\lim_{m \rightarrow \infty} |f_t(z) - f_{\ell^m(t)}^m(z)| = 0.$$

$$\lim_{m \rightarrow \infty} \mathbb{P}(N_{r, \ell^m(t)}^m \leq 1) = 1, \quad r = 1, \dots, m,$$

$$\forall M > 0, \quad \lim_{m \rightarrow \infty} \mathbb{P}(N_{\ell^m(t)}^m > M) = 0.$$

$$C_M := \{N_{\ell^m(t)}^m > M\} \bigcap_{r=1}^m \{N_{r, \ell^m(t)}^m \leq 1\}.$$

$$\forall M > 0, \quad \lim_{m \rightarrow \infty} \mathbb{P}(C_M) = 0.$$

$$\forall \varepsilon > 0, \quad \lim_{m \rightarrow \infty} \mathbb{P}\left(|\hat{f}_{\ell^m(t)}^m(z) - f_{\ell^m(t)}^m(z)| < \varepsilon\right) = 1.$$

FLIGHT MECHANICS MODELS

$$\rho = \frac{P}{R_s SAT}$$

$$SAT(h) = T_0 + \alpha_T h, \quad SAT(TAT, M) = \frac{TAT}{1 + \frac{\lambda-1}{2} M^2}$$

$$M = \frac{V}{V_{sound}} = \frac{V}{(\lambda R_s SAT)^{\frac{1}{2}}}$$

CONSUMPTION X ACCEPTABILITY TRADE-OFF

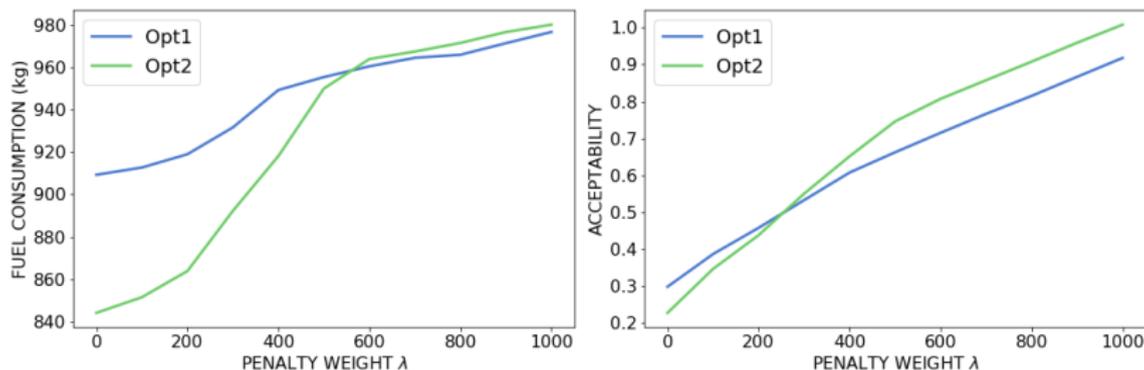
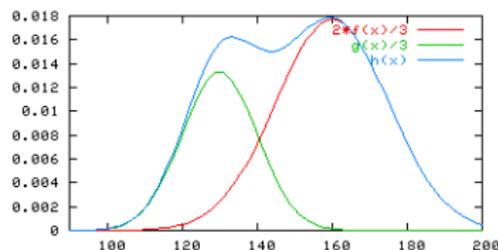


FIGURE: Average over 20 flights of the fuel consumption and MML score (called acceptability here) of optimized trajectories with varying MML-penalty weight λ .

GAUSSIAN MIXTURE MODEL FOR MARGINAL DENSITIES

$$f_t(z) = \sum_{k=1}^K w_{t,k} \phi(z, \mu_{t,k}, \Sigma_{t,k}),$$



$$\sum_{k=1}^K w_{t,k} = 1, \quad w_{t,k} \geq 0,$$

$$\phi(z, \mu, \Sigma) := \frac{1}{\sqrt{(2\pi)^d \det \Sigma}} e^{-\frac{1}{2}(z-\mu)^\top \Sigma^{-1}(z-\mu)}.$$

Assuming that the number of components is known, the weights $w_{t,k}$, means $\mu_{t,k}$ and covariance matrices $\Sigma_{t,k}$ need to be estimated.

MAXIMUM LIKELIHOOD PARAMETERS ESTIMATION

For $K = 1$, maximum likelihood estimates have closed form:

$$\mathcal{L}(\mu_{t,1}, \Sigma_{t,1} | z_1, \dots, z_N) = \prod_{i=1}^N \frac{1}{\sqrt{(2\pi)^d \det \Sigma_{t,1}}} e^{-\frac{1}{2}(z_i - \mu_{t,1})^\top \Sigma_{t,1}^{-1} (z_i - \mu_{t,1})}$$

$$\hat{\theta} := (\hat{\mu}_{t,1}, \hat{\Sigma}_{t,1}) = \arg \min_{(\mu_{t,1}, \Sigma_{t,1})} \sum_{i=1}^N \left(\log \det \Sigma_{t,1} + (z_i - \mu_{t,1})^\top \Sigma_{t,1}^{-1} (z_i - \mu_{t,1}) \right)$$

$$\hat{\mu}_{t,1} = \frac{1}{N} \sum_{i=1}^N z_i, \quad \hat{\Sigma}_{t,1} = \frac{1}{N} \sum_{i=1}^N (z_i - \hat{\mu}_{t,1})(z_i - \hat{\mu}_{t,1})^\top.$$

EM ALGORITHM

- Hidden random variable J valued on $\{1, \dots, K\}$,
- If i^{th} observation $J_i = k$, then z_i was drawn from the k^{th} component,
- Group observations by component and compute $(\hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})$ with $K = 1$ maximum likelihood formulas.

EXPECTATION-MAXIMIZATION - [DEMPSTER ET AL., 1977]

Initialization: $\hat{\theta} = (\hat{w}_{t,k}, \hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})_{k=1}^K = (w_{t,k}^0, \mu_{t,k}^0, \Sigma_{t,k}^0)_{k=1}^K$,

Expectation: For $k = 1, \dots, K$ and $i = 1, \dots, N$,

$$\hat{w}_{t,k} = \frac{1}{N} \sum_{i=1}^N \hat{\pi}_{k,i}, \quad \hat{\pi}_{k,i} := \mathbb{P}(J_i = k | \hat{\theta}_t, Z_h) = \frac{\hat{\mu}_{t,k} \phi(z_i, \hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})}{\sum_{j=1}^N \hat{w}_{t,k} \phi(z_j, \hat{\mu}_{t,k}, \hat{\Sigma}_{t,k})}.$$

Maximization:

$$\hat{\mu}_{t,k} = \frac{\sum_{i=1}^N \hat{\pi}_{k,i} z_i}{\sum_{i=1}^N \hat{\pi}_{k,i}},$$

$$\hat{\Sigma}_{t,k} = \frac{\sum_{i=1}^N \hat{\pi}_{k,i} (z_i - \hat{\mu}_{t,k})(z_i - \hat{\mu}_{t,k})^T}{\sum_{i=1}^N \hat{\pi}_{k,i}}.$$